The province of Alberta, Canada, is transforming its health care system, moving from a focus on acute care services to an emphasis on care in the community that meets the health and social needs of its population. An important aspect of community-based care is the patient’s medical home, where most care is led by a consistent primary care provider. This concept, known as continuity of care, was highlighted decades ago, and mechanisms to support information exchange and disease management within an interpersonal relationship make continuity of care a tenet of primary care. Studies have shown that patients with high continuity of care with a single provider have better outcomes, such as fewer hospital admissions and emergency department encounters, improved delivery of preventive care services, increased adherence to medications and enhanced satisfaction, as well as lower costs. Access to primary care is about a patient’s opportunity to receive timely, appropriate and quality health care services.

It is widely agreed that the opportunity for a patient to receive care when needed is associated with better patient and system outcomes. To our knowledge, the question of precisely how delaying primary care intersects with continuity of care has not been examined empirically. Bennett postulated that continuity of care and access to primary care were not independent concepts. Intuitively, if the delay in access to one’s own physician is too great from the patient’s perspective, it follows that the patient may seek care from another clinic or at the emergency department. Improving access by reducing the delay in obtaining an appointment with one’s primary care physician may be one mechanism to improve continuity of care.

Association between continuity and access in primary care: a retrospective cohort study

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Abstract

Background: Continuity of care is a tenet of primary care. Our objective was to explore the relation between a change in access to a primary care physician and continuity of care.

Methods: We conducted a retrospective cohort study among physicians in a primary care network in southwest Alberta who measured access consistently between 2009 and 2016. We used time to the third next available appointment as a measure of access to physicians. We calculated the provider and clinic continuity, discontinuity and emergency department use based on the physicians’ own panels. Physicians who improved, worsened or maintained their level of access within a given year were assessed in multilevel models to determine the association with continuity of care at the physician and clinic levels and the emergency department.

Results: We analyzed data from 190 primary care physicians. Physicians with improved access increased provider continuity by 6.8% per year, reduced discontinuity by 2.1% per year, and decreased emergency department encounters by 78 visits per 1000 patients per year compared to physicians with stable access. Physicians with worsening access had a 6.2% decrease in provider continuity and an increased number of emergency department encounters (64 visits per 1000 panelled patients per year) compared to physicians with stable access.

Interpretation: Changes in access to primary care can affect whether patients seek care from their own physician, from another clinic or at the emergency department. Improving access by reducing the delay in obtaining an appointment with one’s primary care physician may be one mechanism to improve continuity of care.
shown that when timely access to primary care is mandated by government bodies, continuity of care with the patient’s physician decreases.31

The connection between primary care access (i.e., the length of time a patient waits for an appointment with his or her physician) and continuity of care is complex. We designed this study to explore the relation between the 2 concepts by selecting a group of physicians who measured access consistently over a period of 8 years and determining where their patients sought care when access changed.

Methods

Study design and setting
We conducted a retrospective observational cohort study to explore the association between primary care access and continuity of care outcomes of 205 primary care physicians participating in the Chinook Primary Care Network in Alberta from 2009 to 2016. The Chinook Primary Care Network, which was established in 2005, serves an urban centre of 100 000 and 12 small rural communities in southwest Alberta. Primary care networks in Alberta are the common model of primary care delivery consisting of physicians and allied health care professionals delivering care in their communities.32 The Chinook Primary Care Network was an early adopter of office practice design concepts,33 learning to measure delay for physician appointments and balancing patient demand with physician supply; data were collected regularly for improvement purposes. Additional information about the Chinook Primary Care Network and its adoption of office practice redesign concepts can be found in Appendix 1 (available at www.cmajopen.ca/content/8/4/E722/suppl/DC1).

Study cohort
For each year of the study, network physicians were included in the cohort if they had a physician panel and at least 13 measures of access. A physician panel is a list of patients for whom the physician was the most responsible provider of continuous and longitudinal primary health care. Physicians reviewed their panels annually and reconfirmed their most responsible provider status with their patients (e.g., attachment).34 We extracted panels from 2009 to 2016 from the physician’s electronic medical record, which represented the panel of the physician for a specific year (e.g., the physician panel for 2011 was extracted from the electronic medical record on Jan. 1, 2012, and used to assess the activity of that physician’s patients for 2011). We measured primary care access using the third next available (TNA) appointment35 metric, which is the delay to get an appointment (typically 10–15 min in duration, for routine primary care encounters). A TNA value of 0 indicates that a patient could have a same-day appointment, whereas a TNA value of 14 indicates that a patient would experience a 2-week delay to get an appointment. Third next available appointments for short appointments were recorded in an online database at the same time on Tuesdays by clinic staff; therefore, a week was defined as starting on Tuesday and extending to the following Monday. We recorded the TNA measures for each physician between Jan. 5, 2009, and Jan. 2, 2017. Thirteen TNA measures represented at least 1 measure of TNA appointments each month. Most physicians had more than 43 weekly TNA measures each year.

Outcome measures
Our goal was to associate the weekly TNA value throughout a year directly with the activity of the panelled patients. Therefore, we matched each physician panel to the Alberta Health Practitioner Claims database to determine the number of visits these panelled patients made to a clinic each week. From our previous example, the physician panels for 2011 were matched to the primary care encounters that occurred in 2011. This step ensured that we could link the activity of the physician’s panel directly to the TNA measure on a weekly basis. Patient activity was linked to 1 encounter rather than multiple claims; therefore, we defined 1 visit as the same patient seeing the same physician in the same clinic on the same day. We tabulated the weekly number of visits the panelled patients made to their physician (using the same definition of “week” as used for the TNA metric) and divided it by the total number of weekly primary care visits. This calculation equalled the provider continuity and is based on the Known Provider Continuity Index.36

We also calculated a weekly clinic continuity outcome, taking the weekly number of visits the panelled patients made to other physicians within the clinic and dividing it by the total weekly primary care visits. We further calculated a weekly discontinuity outcome by taking the number of visits the panelled patients made to a physician at another clinic and dividing it by the total number of primary care visits. We matched the clinics to the Alberta Health Services Distance and Drive-Time Look-Up map and included only visits to clinics that were within a 50-km driving radius of where the panelled physician was located.37 More detailed methodology on the continuity measures can be found in Appendix 1.

We also linked the panels for each physician to the National Ambulatory Care Reporting System, which is used in Canada for collecting and reporting on all levels of ambulatory care including that provided at emergency departments. Only nonscheduled visits to emergency departments and urgent care centres were included. We defined 1 emergency department encounter as a visit by a panelled patient to an emergency department within a 50-km driving radius of the panelled physician’s location.37 We then calculated a weekly emergency department encounter outcome, using the same definition of week described above.

Confounding variables
We selected physician practice and panel variables that have been shown to influence appointment delay and continuity6,38–47 (Table 1). We also calculated the starting TNA value as a confounding variable to assess whether having a higher starting value influenced the outcome of interest by taking the average of the first 3 TNA measures.
Statistical analysis
We assessed physicians’ weekly appointment delay and continuity outcomes in 1-year segments (January to December). We conducted a series of linear regressions to identify physicians with appointment delays that improved (statistically significant negative trajectory, \( p < 0.05 \)), worsened (statistically significant positive trajectory, \( p < 0.05 \)) or remained stable (not statistically significant, \( p \geq 0.05 \)) for each year of the study. Grouping the physicians in this manner created 3 TNA trajectory categories (improved, worsened, stable).

We used nested multilevel mixed-effects regression models to assess the relation between the continuity of care outcomes (provider and clinic continuity, discontinuity and emergency department visits) and annual TNA changes, and adjusted for potential confounding patient and physician factors. We examined physician and clinic variations using multilevel mixed-effect regression models with physician nested by year and week. We determined components of variation in the multilevel model by intraclass correlation, which, in our study, meant examining the ratio of variance within each level of our model (weekly TNA measures, physicians and year).

We examined the association between TNA trajectory categories and the weekly continuity trajectory using a group-by-time interaction. The model included a random intercept for each physician within a given year. In this way, our study design allowed us to estimate how continuity changed when physicians’ appointment delays improved, worsened or stayed the same, after adjusting for relevant confounders and accounting for the association among observations within physicians in a given year. A visual representation of the model is provided in Figure 1. We created scatterplots for each TNA trajectory category to visualize the relation with the continuity of care outcomes without adjustment of any factors. This step satisfied our assumption of common support.

### Table 1: Practice- and panel-level confounding variables

<table>
<thead>
<tr>
<th>Practice</th>
<th>Panel</th>
</tr>
</thead>
<tbody>
<tr>
<td>No. of physicians in the clinic</td>
<td>Physician panel size</td>
</tr>
<tr>
<td>Gender of physicians</td>
<td>Age of panel</td>
</tr>
<tr>
<td>Location of clinic</td>
<td>No. of patients on multiple panels</td>
</tr>
<tr>
<td>No. of days worked per week</td>
<td>Panel complexity</td>
</tr>
<tr>
<td>–</td>
<td>No. of female patients</td>
</tr>
</tbody>
</table>

*The Canadian Institute of Health Information’s Population Grouper was used to infer patient complexity.*

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**Figure 1:** Multilevel model focused on primary care physicians (PCPs), measuring weekly access each year of the study. Week of the access measure was nested in the year for each PCP (example shown for 1 physician), with adjustment for potential confounding patient and physician factors.
We performed all statistical analysis using Stata version 13.1 (StataCorp).

**Ethics approval**
This study was approved by the University of Alberta Health Research Ethics Board 3.

**Results**
Of the 205 physicians, 190 met our inclusion criteria (≥ 13 weekly measures of TNA in a calendar year) (Table 2). The number of physicians increased annually between 2009 and 2016 (from 81 to 133), which led to a corresponding increase in the number of panelled patients each year (from 110 868 to 169 653). The number of clinics and the number of physicians per clinic increased over the first 3 years of the study, whereas the proportion of female physicians remained stable, and the proportion of rural physicians decreased. The panel size, mean age, complexity, female composition and representation of older patients all remained consistent over the study period. There was a decrease in the overall proportion of patients present on multiple physician panels (from 15 410 [13.9%] to 19 330 [11.4%]) (Table 2).

Overall, the 4 outcome variables appeared relatively stable year upon year when viewed at the aggregate level (Table 3). The majority of physicians had appointment delay and continuity outcomes for more than 1 year during the 8-year study period. Of the 872 annual TNA trajectories, 96 (11.0%) improved, 669 (76.7%) remained stable and 107 (12.3%) worsened. We combined weeks for each year of the study period to depict each of the average weekly outcome measures segmented by the TNA exposure groups (Figure 2). We then applied a linear regression line to each scatterplot.

For physicians with an improved annual TNA trajectory, provider continuity ($\beta = 0.21$, $p < 0.001$, $R^2 = 34.0\%$) and clinic continuity ($\beta = 0.06$, $p < 0.001$, $R^2 = 35.1\%$) increased.

### Table 2: Descriptive statistics for confounding variables for each study year

<table>
<thead>
<tr>
<th>Variable</th>
<th>Year: mean ± SD*</th>
</tr>
</thead>
<tbody>
<tr>
<td>No. of physicians with a panel</td>
<td>86</td>
</tr>
<tr>
<td>No. of physicians excluded (&lt; 13 TNA measures within the year)</td>
<td>5</td>
</tr>
<tr>
<td>No. of physicians included in analysis</td>
<td>81</td>
</tr>
<tr>
<td>Confounding variables</td>
<td></td>
</tr>
<tr>
<td>Practice characteristics</td>
<td></td>
</tr>
<tr>
<td>No. of physicians</td>
<td>81</td>
</tr>
<tr>
<td>No. of clinics</td>
<td>17</td>
</tr>
<tr>
<td>Female physicians, no. (%)</td>
<td>21 (25.9)</td>
</tr>
<tr>
<td>Rural physicians, no. (%)</td>
<td>47 (58.0)</td>
</tr>
<tr>
<td>No. of days worked per week</td>
<td>4.0 ± 1.5</td>
</tr>
<tr>
<td>No. of physicians per clinic</td>
<td>7.1 ± 3.4</td>
</tr>
<tr>
<td>Panel characteristics</td>
<td></td>
</tr>
<tr>
<td>No. of panelled patients</td>
<td>110 868</td>
</tr>
<tr>
<td>Physician panel size</td>
<td>1458.8 ± 779</td>
</tr>
<tr>
<td>Age of patients, yr</td>
<td>39.6 ± 6.8</td>
</tr>
<tr>
<td>Patients on multiple panels, %</td>
<td>13.9 ± 6.9</td>
</tr>
<tr>
<td>Complex patients, %</td>
<td>5.2 ± 1.8</td>
</tr>
<tr>
<td>Female patients, %</td>
<td>55.2 ± 13.5</td>
</tr>
<tr>
<td>Patients aged &gt; 60 yr, %</td>
<td>213.3 ± 9.7</td>
</tr>
</tbody>
</table>

Note: SD = standard deviation, TNA = third next available appointment.
*Except where noted otherwise.
Table 3: Means of the outcome variables for each year of the study period

<table>
<thead>
<tr>
<th>Outcome variable</th>
<th>Year; mean ± SD</th>
</tr>
</thead>
<tbody>
<tr>
<td>Provider continuity</td>
<td>2009: 58.7 ± 21.9</td>
</tr>
<tr>
<td>All-cause emergency department visits ≤ 50 km per 1000</td>
<td>2009: 1.14 ± 0.6</td>
</tr>
</tbody>
</table>

Note: SD = standard deviation.

Figure 2: Scatterplot and linear regression of average weekly outcome measures within each third next available appointment exposure group. Note: CC = clinic continuity, DC = discontinuity, ED = emergency department, PC = provider continuity.
whereas discontinuity ($\beta = -0.07, p < 0.001, R^2 = 51.0\%$) and emergency department visit rates ($\beta = -0.002, p < 0.001, R^2 = 29.3\%$) decreased (Figure 2).

When the annual TNA trajectories worsened, provider continuity ($\beta = -0.22, p < 0.001, R^2 = 41.8\%$) and clinic continuity ($\beta = -0.05, p < 0.001, R^2 = 37.5\%$) decreased, whereas discontinuity ($\beta = 0.01, p < 0.001, R^2 = 10.6\%$) and emergency department visit rates ($\beta = 0.002, p < 0.001, R^2 = 30.5\%$) increased.

When the TNA trajectories remained stable, provider continuity ($\beta = 0.02, p = 0.4, R^2 = 1.4\%$) did not change; clinic continuity ($\beta = 0.01, p < 0.05, R^2 = 10.4\%$) and discontinuity ($\beta = 13.4, p < 0.001, R^2 = 45.6\%$) increased, and the emergency department visit rate decreased ($\beta = -0.002, p < 0.001, R^2 = 35.7\%$).

With the exception of the starting TNA variable, the confounding variables were balanced across the 3 TNA exposure groups (Table 4). We stratified the starting TNA values into 3 groups (< 5 d, 5–10 d and > 10 d) to explore the impact on the outcome variables and found it affected the degree of change in continuity over time but not the direction of effect.

All confounding variables were included in each of the 4 separate multilevel regression models, 1 for each outcome of interest. The key output within each was the difference in adjusted outcome trajectories during 1 year when TNA improved or worsened compared to when TNA was stable. The $\beta$ coefficients and 95% confidence intervals for each outcome variable by TNA exposure are presented in Table 5. The full model outputs, including the intraclass correlation, are summarized in Appendix 1, Supplemental Table S3.

Physicians who improved their TNA over a 1-year period achieved improvements in provider continuity, discontinuity and emergency department use by their panelled patients as compared to physicians with stable TNA. They saw an improvement in provider continuity of 6.8% (0.13 × 52 wk) per year ($p < 0.001$), reduced discontinuity of 2.1% (0.04 × 52 wk) per year and fewer emergency department visits, by 78 visits per 1000 panelled patients per year (1.5 × 52 wk) ($p < 0.05$) (Table 5). There was no change in clinic continuity ($p = 0.2$).

In the group of physicians among whom TNA worsened over the year, provider continuity decreased by 6.2% (–0.12 × 52 wk) per year as compared to physicians whose TNA was stable ($p < 0.001$) (Table 5). There was no change in clinic continuity ($p = 0.2$) or discontinuity ($p = 0.4$). Emergency department visits increased by 64 visits per 1000 panelled patients per year (1.2 × 52 wk) ($p < 0.1$) compared to physicians with stable TNA.

**Interpretation**

When physicians improved their appointment delay, provider continuity increased, patients’ attendance with external providers was lower, and patients’ use of the emergency department decreased. The opposite result — decreased provider continuity and increased emergency department use — was found when physicians worsened their availability to their patients.

Our findings support the following observations on the impact of appointment delay on continuity of care. When faced with a delay for an appointment, patients chose to break their continuity of care with their attached physician and seek care at another clinic or in the emergency department. Although we are unable to claim a causal link between

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### Table 4: Confounding variables within the third next available appointment exposure groups

<table>
<thead>
<tr>
<th>Variable</th>
<th>Improved n = 96</th>
<th>Stable n = 669</th>
<th>Worsened n = 107</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Practice characteristics</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Female physicians, no. (%)</td>
<td>38 (39.6)</td>
<td>186 (27.8)</td>
<td>37 (34.6)</td>
</tr>
<tr>
<td>Rural physicians, no. (%)</td>
<td>49 (51.0)</td>
<td>338 (50.5)</td>
<td>51 (47.7)</td>
</tr>
<tr>
<td>No. of days worked per week</td>
<td>4.1 ± 1.4</td>
<td>4.0 ± 1.4</td>
<td>4.0 ± 1.4</td>
</tr>
<tr>
<td>No. of physicians per clinic</td>
<td>8.8 ± 4.3</td>
<td>8.0 ± 4.0</td>
<td>8.0 ± 3.9</td>
</tr>
<tr>
<td><strong>Panel characteristics</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Physician panel size</td>
<td>1412.8 ± 934.4</td>
<td>1388.1 ± 791.7</td>
<td>1363.3 ± 938.2</td>
</tr>
<tr>
<td>Age of panel, yr</td>
<td>39.2 ± 5.3</td>
<td>40.0 ± 6.6</td>
<td>39.2 ± 6.9</td>
</tr>
<tr>
<td>Patients on multiple panels, %</td>
<td>13.3 ± 7.8</td>
<td>13.1 ± 8.4</td>
<td>12.4 ± 8.4</td>
</tr>
<tr>
<td>Complex patients, %</td>
<td>5.6 ± 2.4</td>
<td>5.3 ± 2.0</td>
<td>5.4 ± 1.8</td>
</tr>
<tr>
<td>Female patients, %</td>
<td>57.4 ± 12.8</td>
<td>54.2 ± 13.5</td>
<td>55.7 ± 13.7</td>
</tr>
<tr>
<td>Starting TNA value</td>
<td>11.7 ± 10.7</td>
<td>4.4 ± 5.0</td>
<td>4.1 ± 4.1</td>
</tr>
</tbody>
</table>

Note: SD = standard deviation, TNA = third next available appointment.

*Except where noted otherwise.
The office practice redesign philosophy operant in the Chinook Primary Care Network encourages each physician to “take care of your own,” so the option for patients to see another physician in the same clinic was typically not available. Contingency plans when the attached physician was away for more than a few days (e.g., on vacation) may include offering appointments with another physician practising in the clinic, but those situations reflect an exception.

Our study confirms a long-held assumption that continuity of care and access are not independent concepts. Some jurisdictions have reported unintended consequences on continuity of care when access is given precedence; we assume that the opposite may be true when continuity of care, especially relational continuity of care, is given precedence over appropriate access. In this era of public expectations of instant access, there has been a proliferation of walk-in style clinics and smartphone apps that threaten continuity of care. Primary care access is a critical consideration in the provision of high-quality and effective health care; however, its balance with continuity of care should not be overlooked.

Physicians cannot be available at all times, so it is reasonable to posit that physician-led team-based care in the patient’s medical home is a plausible strategy to increase access to a consistent team and, subsequently, continuity of care. Metrics that can assess “team continuity” should be developed. The culture and funding of primary care needs to best meet the presenting need.

In this study, we focused on identifying physicians who improved, worsened or maintained their level of access within a given year; we did not address the rate of change or the different levels of access within the stable group (e.g., physicians who maintained same-day or next-day access throughout the year compared to those who maintained ≥ 20-d access). Our team plans to conduct studies to address these topics and attempt to determine the optimal access threshold within primary care.
Limitations
A limitation of this study was its restriction to 1 primary care network that does not serve a large metropolitan centre where more service options (e.g., walk-in clinics) may be available. The observed relation between TNA appointment and provider continuity will undoubtedly be more complex in the presence of other care alternatives. However, our findings are still generalizable to other parts of Alberta and to other Canadian provinces; if access to the physician is appropriate from the patient’s perspective, provider continuity will be maintained. The longitudinal measurement of TNA and the data generated by using annual panels allowed us to explore this relation in greater detail.

It is unclear whether the increased emergency department use observed may equate to a deteriorating condition; rather, it may reflect a convenient source of primary care.

The TNA metric we chose to infer access may also be a limitation of our study. We recognize that access to primary care is a large concept that includes both having a physician and getting access to that physician’s services. The TNA metric is limited to the delay the patient experiences when attempting to book an appointment with his or her physician. Since its value can be obtained only by viewing the physician’s schedule at a specific day and time each week, there is no way to verify its accuracy. However, in our study, TNA appointments were measured for quality-improvement purposes, and because no TNA targets were mandated, they were not likely to be gamed; we are thus reasonably confident that the TNA measures used are a valid inference of delay. Nevertheless, other jurisdictions have failed to implement TNA measurement successfully. We concede that implementing TNA measurement would be difficult if the concept of TNA was introduced without the support of an office practice redesign learning collaborative.

We used statistical significance and trajectory direction to define our TNA exposure groups to decrease the uncertainty regarding an annual change. If we had relied only on the steepness of the trajectory, we may have run the risk of classifying TNA as improving or worsening inappropriately. Within the stable TNA trajectory group, most physicians had a $p > 0.2$. The certainty around appropriate physician classification in the exposure category would have been low for many of these slopes if we had not taken into account the $p$ value.

Conclusion
Our findings suggest that changing appointment delay in primary care can influence how patients choose to use the health care system. Furthermore, it can affect provider continuity, discontinuity and emergency department use, which, in turn, can affect health and system outcomes. As Alberta and other jurisdictions reform their health care systems to ensure patients receive appropriate care in the community, focusing on reducing delay in obtaining appointments with physicians practising in the community should be considered as a focal point of primary care form. However, improving access to primary care should not be done at the expense of continuity of care.

References
47. CIHI’s population grouping methodology 1.0. Ottawa: Canadian Institute for Health Information; 2016.

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Contributors: Lisa Cook, Richard Golonka, Charles Cook and Scott Oddie were responsible for the conception and design of the study. Lisa Cook, Richard Golonka, Charles Cook, Robin Walker, Peter Faris, Rebecca Love and Scott Oddie were involved in data analysis and interpretation. Lisa Cook drafted the manuscript, and Richard Golonka, Charles Cook, Robin Walker, Peter Faris, Shannon Spenceley, Richard Lewanczuk, Robert Wedel, Rebecca Love, Cheryl Andres, Susan Byers, Tim Collins and Scott Oddie revised it for important intellectual content. All of the authors approved the final version to be published and agreed to be accountable for the work.

Data sharing: The data are not available for use by other researchers.

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