Opioid use in North America has continued to increase dramatically over much of the past decade.1,2 Canada ranks with the United States as having the highest prescription opioid consumption in the world.3 Several US studies of health claims have found that opioid use in pregnancy has also increased over time.4–7 Recent studies have reported that 14% to 22% of women filled at least 1 prescription for an opioid during pregnancy, with exposures as high as 42% in some US states.5,6,8 Although the effects of opioid use on the developing fetus are poorly understood and findings are varied, associations with complications, including neural tube defects and cardiac septal defects, have been found.9–11 Neonatal abstinence syndrome, a drug-withdrawal syndrome that most commonly occurs after in utero exposure to opioids, has become one of the fastest growing reasons for neonatal hospital admissions.11,12 Recent studies of opioid exposure in pregnancy have reported incidences of 4.3 to 5.9 cases of neonatal abstinence syndrome per 1000 births.11,14

Considering the overall population use of opioids in North America and the potential harms of opioid exposure during pregnancy to the developing fetus and the neonate, it is important to study prescribing patterns in order to understand the potential implications from a public health perspective. Two large retrospective analyses of outcomes related to antenatal and postpartum opioid use have been performed in Ontario, Canada, in addition to small cohort studies evaluating the incidence of narcotic abuse during pregnancy in communities in northwestern Ontario.13,15–17 The objectives of this study were to determine, for the period 2001 to 2013, opioid use before and during pregnancy in a complete population-based cohort.
the frequency of use and volume of consumption of opioids among pregnant women in Manitoba, Canada, before and during pregnancy and to determine the rate of continued use of opioid agents through each trimester.

Methods

Setting and data sources
We performed an observational study of pregnancy and opioid use using administrative health care claims data from the Population Health Research Data Repository, housed at the Manitoba Centre for Health Policy. We used the Medical Services database, which captures all fee-for-service medical claims, to determine the date of first pregnancy-related physician visit. We also used the Hospital Discharge Abstract database, which contains summary data for each hospital stay; the Manitoba Health Registry, which contains basic demographic information on persons registered with Manitoba Health; and the Drug Program Information Network, for prescription drug claims. We used neighbourhood income quintiles from public-use census files to determine socioeconomic status. These data sets are linked at the person level with an encrypted identifier and have been used extensively in health services research.

Design
We created a cohort of all pregnant women in Manitoba with resulting live birth, stillbirth or intratuerine death who were registered with Manitoba Health during the pregnancy, in the year before the pregnancy and 3 months after the delivery. We identified pregnancies by codes from the International Classification of Diseases, 9th revision (V27.0–V27.9, 656.4) and corresponding codes from the International Statistical Classification of Diseases and Related Health Problems, 10th revision (Z37.0-Z37.9, O36.4). We included women with more than 1 pregnancy during the study period as multiple pregnancy observations, and excluded molar and ectopic pregnancies and pregnancies ending in abortion. Because of uncertainty about delivery date, we excluded women with length of stay in hospital more than 7 days and no newborn data that would allow assignment of a birth date. Gestational age is available through hospital discharge data (maternal and newborn records) for 98.6% of live births, stillbirths and intratuerine deaths. We determined date of conception by subtracting gestational age at delivery from the maternal hospital admission date.

We considered the pre-pregnancy period (3 mo before conception) and each trimester of pregnancy. Given that the exposure per person may vary (because not all women complete all time periods), we determined a weighted number of pregnancies in each period. We summed the actual number of days each woman contributed during a given period for all pregnancies, and then divided the total number of days that all pregnancies contributed to a period by the possible number of days in that period (e.g., 91 d for each trimester). Inclusion in the opioid user cohort was defined by filling of prescriptions for opioids at any time after date of conception. We analyzed this cohort for patterns of use over time starting 3 months before pregnancy and through each trimester. To evaluate opioid use by women who would be certain they were pregnant, we performed a comparative analysis in which inclusion in the user cohort was defined by filling of prescriptions for opioids after the first pregnancy-related visit. In-hospital medication use was not available and was therefore not included in the study.

For each opioid prescription, we calculated the total amount dispensed by multiplying the total number of tablets (or total volume of liquid) by the strength of the pills (or concentration of the liquid). We then converted the calculated total opioid (in milligrams) to oral morphine equivalents (MEQ) using the morphine equivalence ratios of the Canadian Guideline for Safe and Effective Use of Opioids for Chronic Non-Cancer Pain, with supplementary references used for opioids not listed in the Canadian guideline. We excluded methadone from the analysis because most methadone use was in the form of compounded products that did not allow reliable determination of the actual quantity of medication.

Statistical analysis
We used \( \chi^2 \) and \( t \) tests to compare opioid use at the beginning and end of the study period. To evaluate trends in rates of use over the study period, we ran Poisson regression models with log of the population as the offset. To evaluate opioid use across pregnancy trimesters, we ran logistic regression generalized estimating equation models; these models accounted for the correlated observations (i.e., repeated measures) of individuals across the pregnancy. All analyses were performed with SAS version 9.3 (SAS Institute Inc.).

Ethics approval
Approval for this study was obtained from the University of Manitoba Health Research Ethics Board (HREB HS16618 [H2013:332]).

Results
During the study period, there were a total of 245 704 pregnancies. There were 49 914 exclusions for molar and ectopic pregnancies and abortions, and 5633 exclusions related primarily to multiple pregnancy outcomes per event date and non-availability of gestational age. Women not having continuous registration in Manitoba for health coverage accounted for an additional 15 309 exclusions. After these exclusions, there were 174 848 completed pregnancies eligible for analysis (173 680 live births and 1168 stillbirths and intratuerine deaths), which represented a weighted value of 175 174 pregnancies.

Over the study period, an opioid was dispensed for 13 408 women (7.7%) during their pregnancy. The median age was 28 years for the overall cohort and 27 years for the cohort with exposure to opioids. The incident pregnancy was the first pregnancy for 29.3% of women in the overall cohort and for 18.5% of those with exposure to opioids. The lowest income quintile accounted for 26.3% of the overall cohort, compared with 40.0% of those with exposure to opioids (Table 1).
Over the study period, there was a modest increase in the proportion of women who filled prescriptions for opioids at any time during pregnancy, from 7.3% in 2001 to 7.7% in 2013 ($p = 0.03$) (Figure 1). In the 3 months before pregnancy, an opioid was dispensed for 11 692 women (6.7%). Opioid use fell to 4.2% during the first trimester of pregnancy and further declined to 3.0% in the second and 2.9% in the third trimester ($p < 0.001$ for each trimester compared with the 3 mo before pregnancy and for second and third trimesters compared with the first trimester). A similar pattern of usage decline from the pre-pregnancy period through the trimesters was observed from the start to the end of the study period (Figure 1).

There was a significant increase in consumption of prescription opioids across the study period. For women who filled prescriptions for opioids at any time during pregnancy, the mean MEQ per pregnancy increased more than 4-fold over the study period, from 284 mg in 2001 to 1218 mg in 2013 ($p < 0.001$) (Figure 2). The mean MEQ over the 3 months before pregnancy was 421 mg, similar to the 493 mg, 545 mg and 511 mg used in the first, second and third trimesters of pregnancy, respectively ($p > 0.9$ for all comparisons). A similar pattern from the pre-pregnancy period through the trimesters was observed from the start to the end of the study period (Figure 2).

### Table 1: Characteristics of the study population for completed pregnancy cohort

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Group; weighted % of women*</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Total cohort $n = 175 174$</td>
</tr>
<tr>
<td>Pregnancy type</td>
<td></td>
</tr>
<tr>
<td>Live birth</td>
<td>99.3</td>
</tr>
<tr>
<td>Stillbirth or intrauterine death</td>
<td>0.7</td>
</tr>
<tr>
<td>Age at delivery, yr, median (IQR)</td>
<td>28 (23–32)</td>
</tr>
<tr>
<td>Age at delivery, yr</td>
<td></td>
</tr>
<tr>
<td>≤ 19</td>
<td>9.0</td>
</tr>
<tr>
<td>20–24</td>
<td>21.3</td>
</tr>
<tr>
<td>25–29</td>
<td>29.7</td>
</tr>
<tr>
<td>30–34</td>
<td>26.5</td>
</tr>
<tr>
<td>≥ 35</td>
<td>13.6</td>
</tr>
<tr>
<td>No. of past pregnancies, including current pregnancy</td>
<td></td>
</tr>
<tr>
<td>1</td>
<td>29.3</td>
</tr>
<tr>
<td>2</td>
<td>28.6</td>
</tr>
<tr>
<td>3</td>
<td>17.8</td>
</tr>
<tr>
<td>≥ 4</td>
<td>24.2</td>
</tr>
<tr>
<td>No. of past deliveries†</td>
<td></td>
</tr>
<tr>
<td>0</td>
<td>37.6</td>
</tr>
<tr>
<td>1</td>
<td>32.0</td>
</tr>
<tr>
<td>2</td>
<td>15.7</td>
</tr>
<tr>
<td>≥ 3</td>
<td>14.5</td>
</tr>
<tr>
<td>Interval between conception date and first pregnancy-related visit, d, median (IQR)</td>
<td>72 (52–93)</td>
</tr>
<tr>
<td>Income quintile‡</td>
<td></td>
</tr>
<tr>
<td>1 (lowest)</td>
<td>26.3</td>
</tr>
<tr>
<td>2</td>
<td>20.5</td>
</tr>
<tr>
<td>3</td>
<td>18.5</td>
</tr>
<tr>
<td>4</td>
<td>18.4</td>
</tr>
<tr>
<td>5 (highest)</td>
<td>15.9</td>
</tr>
</tbody>
</table>

Note: IQR = interquartile range.
*Except where indicated otherwise.
†Data on number of past deliveries were not available for 0.1% of the total cohort, 0.1% of non-users and 0.1% of users.
‡Data on income were not available for 0.4% of the total cohort, 0.4% of non-users and 0.3% of users.
Figure 1: Opioid use for overall pregnancy cohort over study period.

Figure 2: Mean opioid use, as morphine equivalents (MEQ), for opioid users over the study period. Mean MEQ = total MEQ divided by total weighted number of pregnancies among opioid users.
In the comparative analysis with cohort inclusion defined as filling of any opioid prescription after the first pregnancy-related visit, overall use of opioids was lower than in the primary analysis (4.9%), but a similar pattern of use and consumption was seen over the study period. Use increased modestly from 4.3% in 2001 to 4.8% in 2013 ($p = 0.01$), while mean MEQ increased from 281 mg in 2001 to 1046 mg in 2013 ($p < 0.001$).

Of women in the lowest and highest income quintiles, 11.7% and 5.1% were opioid users, respectively ($p < 0.001$). The mean MEQ during pregnancy did not differ significantly between these 2 quintile groups (773 mg v. 653 mg, $p = 0.5$) (Table 2).

Codeine accounted for most of the opioid prescriptions, having been dispensed to 96.9% of all exposed women and accounting for 66.2% of MEQ. The proportion of the MEQ accounted for by codeine declined over the study period, from 87.8% in 2001 to 56.7% in 2013. There was a corresponding increase in the contribution of higher-potency opioids to the MEQ over this time period (Table 3).

**Interpretation**

Among the more than 175,000 pregnancies in the population of Manitoba, Canada (2001–2013), an opioid was dispensed for 7.7% of women during the pregnancy. Over the study period, the proportion of those who filled prescriptions for opioids in pregnancy increased modestly; however, the amount of prescription opioid used during pregnancy increased dramatically. The proportion of women using prescription opioids declined from the pre-pregnancy period to the first trimester and further declined into the second trimester; however, the volume of opioids used by those who continued opioid use did not decline.

### Table 2: Opioid use and mean MEQ during pregnancy, by income quintile

<table>
<thead>
<tr>
<th>Income quintile</th>
<th>Any time during pregnancy</th>
<th>After first pregnancy-related visit</th>
<th>Overall</th>
<th>Those using opioids after first pregnancy-related visit</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 (lowest)</td>
<td>11.7*</td>
<td>6.9*</td>
<td>773†</td>
<td>689‡</td>
</tr>
<tr>
<td>2</td>
<td>7.2</td>
<td>4.6</td>
<td>720</td>
<td>615</td>
</tr>
<tr>
<td>3</td>
<td>6.5</td>
<td>4.3</td>
<td>543</td>
<td>497</td>
</tr>
<tr>
<td>4</td>
<td>5.8</td>
<td>3.9</td>
<td>562</td>
<td>486</td>
</tr>
<tr>
<td>5 (highest)</td>
<td>5.1*</td>
<td>3.7*</td>
<td>653†</td>
<td>668‡</td>
</tr>
</tbody>
</table>

Note: MEQ = morphine equivalents.  
* $p < 0.001$ for comparison of quintiles 1 and 5.  
† $p = 0.5$ for comparison of quintiles 1 and 5.  
‡ $p = 0.9$ for comparison of quintiles 1 and 5.

### Table 3: Opioids dispensed

<table>
<thead>
<tr>
<th>Opioid</th>
<th>2001</th>
<th>2013</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Use, no. (%) of women*</td>
<td>MEQ, % of total</td>
</tr>
<tr>
<td>Any opioid†</td>
<td>$n = 937$</td>
<td>878</td>
</tr>
<tr>
<td>Codeine</td>
<td>922 (98.4)</td>
<td>878</td>
</tr>
<tr>
<td>Fentanyl</td>
<td>0 (0)</td>
<td>–</td>
</tr>
<tr>
<td>Hydrocodone</td>
<td>S (&lt; 1.0)</td>
<td>&lt; 1.0</td>
</tr>
<tr>
<td>Hydromorphone</td>
<td>S (&lt; 1.0)</td>
<td>&lt; 1.0</td>
</tr>
<tr>
<td>Meperidine</td>
<td>15 (1.6)</td>
<td>4.3</td>
</tr>
<tr>
<td>Morphine</td>
<td>S (&lt; 1.0)</td>
<td>6.6</td>
</tr>
<tr>
<td>Oxycodone</td>
<td>S (&lt; 1.0)</td>
<td>&lt; 1.0</td>
</tr>
<tr>
<td>Tramadol</td>
<td>0 (0)</td>
<td>–</td>
</tr>
</tbody>
</table>

Note: MEQ = morphine equivalents, S = suppressed data (for privacy reasons, data are suppressed for any count greater than 1 but less than 5).  
*Percentages may sum to more than 100% because some patients had multiple opioids dispensed.  
†Buprenorphine, butorphanol and dextropropoxyphene were included in the analysis, but had no counts.
Codeine was used by most of these women, and use of this drug showed a small decline over the study period. The proportion of total MEQ accounted for by codeine decreased significantly over the study period, with a corresponding increase in more powerful opioids. Although we saw a modest increase in the proportion of opioid users over time, more striking was the finding, based on an analysis of MEQ, of a dramatic increase in the overall volume of opioids dispensed; this result suggests that women who used opioids during pregnancy were using more or were switching away from less potent options, such as codeine, to more potent forms, such as morphine and oxycodone, which would make a larger contribution to the MEQ. To our knowledge, no population-based Canadian studies have evaluated the volume of opioids used in pregnancy on the basis of MEQ consumed. Considering that opioids have been associated with neural tube defects, cardiac septal defects and the rise in neonatal hospital admissions associated with neonatal abstinence syndrome, our findings raise some concern.8–14

A previous analysis of the overall Canadian population, for the period 2005–2010, showed a similar increase in opioid use of 13.1%.29 US database studies have also shown increases in opioid use during pregnancy (from the mid-1990s to 2009), but the magnitude of change was generally larger than what we found, with increases in exposure ranging from 0.23-fold to 4.73-fold.4,7 The most up-to-date exposure rates in the US were varied but generally much higher than in Manitoba, ranging from 14% to 42%, with pharmacy dispensing claims data being used to determine opioid exposure.5,6,8

There are several possible reasons for the differences in opioid use between US pharmacy claims studies and our findings. These differences may indicate substantially greater use of opioids by pregnant women in the US. Manitoba has a fully linked drug information network that allows all pharmacists to review any prescription filled in the past and requires a special physician-specific, fraud-resistant prescription for single-entity opioid prescriptions. The variation between studies may also be related to characteristics of the populations studied. Rates of prescription opioid use have been reported to be higher among those enrolled in Medicaid than among commercially insured patients.10 However, Bateman and colleagues, despite having one of the lower estimates among the current US studies, reported a considerably higher rate of prenatal opioid use (14% in a commercially insured US population) than was found in our study. In Manitoba, those with lower economic status have full medication coverage, with minimal or no deductible.11 Thus, our study population, which encompassed all residents in the province, incorporated aspects of both the commercially insured and publicly insured US populations. Our analysis of income quintiles showed that those in the lowest income group had a significantly higher rate of use (11.7%) than those in the highest income group, although still lower than rates found in US studies.

The findings of Handal and colleagues,12 who reported the prevalence of opioid use among pregnant women in Norway between 2004 and 2008 as only 3%, were more similar to our results. Similar to our study, use of codeine accounted for most exposures, at 91%. In addition, codeine use declined from 2.4% to 0.93% from 3 months before pregnancy to the second trimester, a relative reduction of about 60%, which was very similar to that found in our overall pregnancy cohort (from 6.7% to 3.0%). Instead of using MEQ as a marker of volume of use, Handal and colleagues12 assessed divided daily doses (DDD), finding a small reduction from 6.7 DDD in the 3 months before pregnancy to 5 DDD in the second and third trimesters. This result contrasts with our study, using MEQ as the marker, which showed no reduction in volume of use.

Limitations and strengths
Our study had certain limitations. First, determining medication use from pharmacy dispensing claims does not allow assessment of whether the women actually took the medication. In addition, use of pharmacy claims data does not allow capture of use of over-the-counter opioid analgesics that contain low doses of codeine, illicit opioid use or impotent opioid use; as a result, the true extent of opioid use would be underestimated.

Use of the Manitoba Centre for Health Policy database offered several advantages. First, it allowed a Canadian perspective on frequency and utilization patterns of opioid use by pregnant women. Second, data from the Drug Program Information Network accounts for all women, regardless of socioeconomic status, and linkage to socioeconomic status variables allowed for comparisons among socioeconomic strata of the population.

Conclusion
In this study population, many women who used opioids before pregnancy discontinued their use during pregnancy. Although the percentage of opioid users in pregnancy was low and declined across trimesters relative to pre-pregnancy values, the volume of opioids used by those who continued to use these drugs remained constant, relative to the pre-pregnancy period, and increased significantly over time. Given this increasing volume of opioid use over time, these findings indicate a need for continued evaluation of the safety of this group of medications during pregnancy and necessitate the provision of better information regarding the currently understood benefits and risks of pregnancy-related opioid use to guide clinical care.

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Shawn Bugden performed the initial analysis and interpretation of the
data. All of the authors provided comment on refinement of the analysis
and data interpretation. Jamie Falk and Shawn Bugden drafted the
manuscript. All of the authors revised and critically evaluated the
manuscript, approved the final version for publication and agreed to act as
guarantors of the work.

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