

Operative versus nonoperative interventions for common fractures of the clavicle: a meta-analysis of randomized controlled trials

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Abstract

Background: The popularity of surgery for acute displaced midshaft clavicle fractures has been fuelled by early randomized controlled trials (RCTs) showing improved rates of radiographic union and perceived functional benefits compared with nonoperative approaches. We performed a meta-analysis to determine the effect of operative and nonoperative interventions on the risk of secondary operation and complications and on long-term function.

Methods: We search MEDLINE, Embase and the Cochrane Central Register of Controlled Trials for reports of relevant RCTs published to Mar. 7, 2014. Two reviewers assessed eligibility of potential reports and the risk of bias of included trials. The Grading of Recommendations Assessment, Development and Evaluation approach was used to summarize the quality of evidence for all outcomes.

Results: We included 15 RCTs (9 trials comparing operative and nonoperative interventions, 5 comparing implants for operative treatment, and 1 comparing nonoperative treatments). Nonoperative treatments did not differ from operative treatments in the risk of secondary operation (risk ratio [RR] 1.16, 95% confidence interval [CI] 0.58 to 2.35) or all complications (RR 0.90, 95% CI 0.55 to 1.50). One in 4 patients had a complication regardless of the treatment approach. Differences in functional outcomes, although smaller than the threshold for minimal important differences at 1 year, favoured operative interventions (standardized mean difference 0.38, 95% CI 0 to 0.75). Evidence for the type of implant or approach to nonoperative treatment remained inconclusive.

Interpretation: Current evidence does not support the routine use of internal fixation for the treatment of displaced midshaft clavicle fractures. Complication rates were high regardless of the treatment approach.

Clavicle fractures are common injuries, affecting about 22 000 Canadians each year and numbering 1.75 million fractures worldwide.¹⁻⁶ The majority of these fractures are located in the midshaft, accounting for about 80% of all clavicle fractures.^{1,2} Closed midshaft fractures were traditionally treated nonoperatively, a practice largely based on previous studies by Neer and Rowe.^{7,8} In the last decade, evidence challenged the standard of nonoperative treatment, reporting high rates of nonunion (15%–20%), poor early function, and residual sequelae at 6 months following nonoperative management in up to 42% of patients.⁹ Small clinical trials that followed have fuelled a growing popularity to treat these fractures surgically with plates and screws or intramedullary devices; however, these procedures carry inherent surgical risks for infection, implant failure and hardware irritation requiring subsequent removal.^{10,11}

Whether surgery or a conservative approach is the optimal method of management for midshaft clavicle fractures is still an issue of debate. Several trials have compared operative and nonoperative approaches to treatment. In the last 5 years, a number of trials have also investigated various surgical tech-

niques and the use of different implants. Previous reviews focused only on the operative versus nonoperative debate.^{1,9,11,12} Our review adds to this body of literature by providing data from the largest and most recent trial. It also provides a summary of the evidence on surgical techniques for these injuries, as well as nonsurgical options.

We performed a meta-analysis to determine the effect of operative and nonoperative interventions for treating acute displaced midshaft clavicle fractures on the risk of secondary operation and all complications and on long-term function.

Competing interests: Mohit Bhandari has received consultancy fees from Smith & Nephew, Stryker, Amgen, Zimmer, Moximed, Bioventus, Merck, Eli Lilly and Sanofi; he was involved in research funded by Smith & Nephew, DePuy, Eli Lilly, Bioventus, Stryker, Zimmer and Amgen. No other competing interests were declared.

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Methods

We conducted this study according to the methods outlined in the *Cochrane Handbook for Systematic Reviews of Interventions*.¹³ We report our findings in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses (PRISMA) statement.¹⁴

Literature search

We systematically searched MEDLINE (from 1946), Embase (from 1974) and the Cochrane Central Register of Controlled Trials (CENTRAL; from 1948) for articles published up to and including Mar. 7, 2014. Subject headings and subheadings (MeSH terms in MEDLINE and Emtree terms in Embase) were used in various combinations and supplemented with free text (an example of the search strategy is available in Appendix 1, www.cmajopen.ca/content/3/4/E396/suppl/DC1). An RCT filter developed by the Health Information Research Unit at McMaster University¹⁵ was applied to the search. No restrictions on language or publication date were applied. Manual searches of the reference lists of key articles and of “related articles” featured in PubMed were conducted to identify additional articles. We searched conference proceedings (American Academy of Orthopaedic Surgeons, Canadian Orthopaedic Association and Orthopaedic Trauma Association) from the last 5 years and ClinicalTrials.gov to identify relevant unpublished studies.

Study selection

We included RCTs comparing any form of operative or non-operative interventions for acute displaced midshaft clavicle fractures in patients 16 years of age or older. This included studies comparing operative and nonoperative interventions, studies comparing operative implants and studies comparing nonoperative interventions. Two reviewers (T.D. and Y.K.), both with methodologic expertise and one with content expertise, independently screened all titles and abstracts of reports identified through the literature search. Disagreements were carried forward for full-text review. Both reviewers independently reviewed the full text of potentially eligible reports; disagreements were resolved through a consensus process to determine final eligibility.

Data extraction and quality assessment

The same 2 reviewers independently extracted study data using a piloted electronic data collection form. Authors of the included studies were contacted if important data were unclear or not reported. When information was reported by graphical analyses only, the data were derived from the figures using graph digitizing software (GraphClick, Arizona Software).

The primary outcomes for this review were secondary operation, all complications and long-term function (≥ 1 yr). We did not include routine removal of hardware as a secondary operation, because only studies of pin fixation versus nonoperative treatment reported routine removal following fracture healing. This procedure is typically done under local anesthesia with minimal sedation, requires a small incision over the tip of the nail and is not likely to result in complications. Conversely, plate removal

is usually indicated as a result of discomfort and necessitates new admission, surgery under general anesthesia and an additional large incision. Thus, we included only removals that had an indication for removal such as infection, irritation or implant failure. Complications included symptomatic malunion, symptomatic nonunion, loss of primary reduction, hardware irritation, infection, neurologic symptoms, and other issues requiring surgical treatment. The selected complications were chosen because they are considered to be patient-important outcomes or were commonly reported in the identified primary studies.

For the assessment of methodologic quality, both reviewers independently assessed the risk of bias in included trials using the Cochrane Collaboration’s risk-of-bias tool.¹³ They evaluated the quality of evidence in included trials for each outcome using the GRADE approach.¹⁶ Data from RCTs were considered high-quality evidence, but the quality could be rated down because of risk of bias, inconsistency, imprecision, indirectness or publication bias.

Data synthesis

We used risk ratios (RRs) to summarize results for secondary operation and all complications and mean differences to summarize results for functional outcome scores.

We pooled data on secondary operations and all complications from only trials that reported these outcomes; we calculated RRs using the Mantel–Haenszel method and a random-effects model.¹³ The nature and criteria of the primary outcomes selected for our review were such that patients experiencing an event would require surgical intervention or additional medical management. We performed a “none has event” analysis, a variation of “analysis as randomized,”¹⁷ because it is highly plausible that most patients would return for follow-up if unsatisfied with treatment or experienced an adverse event. All patients included at randomization comprised the denominator; those lost to follow-up were assumed to not have had an event.¹⁷ Because some patients lost to follow-up may have experienced an event and sought treatment elsewhere, we performed 2 sensitivity analyses to test the robustness of the assumption made in our primary analysis and investigate the effects of dropouts and exclusions: (a) a complete case analysis and (b) an arm-level assumption analysis, in which the relative incidence among patients with missing data was assigned the same incidence as those followed up in the same study arm.¹⁷

We performed a complete case analysis and used standard mean differences (SMDs) to summarize results for long-term function. The SMDs were weighted according to the inverse variance method and pooled using a random-effects model.^{13,18,19} Minimal important differences (MIDs) were incorporated to aid the interpretation of treatment effects. The MID describes the smallest effect that an informed patient would perceive as beneficial enough to justify a change in management.^{20–24} The MID for the Disabilities of the Arm, Shoulder and Hand (DASH) questionnaire is estimated to be 10.2 points,^{25,26} which was converted to units of standard deviation (SD) using the DASH median SD for each comparison.²⁷ A zone of clinical equivalence based on the converted MID was projected onto the forest plot to aid interpretability of the pooled SMDs.

We quantified heterogeneity using the χ^2 test for heterogeneity and the I^2 statistic.¹³ We developed a priori hypotheses to explain potentially high heterogeneity in treatment effects across trials between intramedullary and plate fixation, between immediate and delayed (1–4 wk) surgical intervention, between 2-fragment and comminuted fractures, and between the presence and absence of selection or detection bias.

We evaluated interobserver agreement for assessments of study eligibility using the Cohen κ coefficient, and we evaluated interobserver agreement for risk-of-bias assessments using weighted κ coefficients;^{28,29} all coefficients were calculated using SPSS software (version 21.0; SPSS Inc.). All tests of significance were 2-tailed, and p values of less than 0.05 were considered significant. To assess for publication bias, we visually inspected a funnel plot for the outcome of long-term function.¹³ The forest plots and funnel plot were generated using Review Manager software (RevMan version 5.2; Nordic Cochrane Centre, Cochrane Collaboration).

Results

Study characteristics

We identified 422 potentially relevant citations through the literature search. Fifteen of these studies proved eligible for inclusion (Figure 1).^{5,10,30–42} The overall agreement between the 2 reviewers for final eligibility was excellent (κ value = 0.94, 95% confidence interval [CI] 0.84–1). All 15 studies were published between 2007 and 2013; their characteristics are summarized in Table 1. Nine studies compared operative with nonoperative treatment.^{5,10,30–36} Two-year follow-up data from one of the trials¹⁰ was reported in a separate publication.⁴³ Five studies compared different operative implants.^{37–41} One placebo-controlled trial managed all fractures nonoperatively.⁴² Twelve studies were reported in English. The other 3 studies^{31,33,38} were translated by reviewers with expertise in systematic review methodology.

The findings from the risk-of-bias assessment are shown in Figure 2. All 15 studies were found to have an uncertain to high overall risk of bias. Blinding of participants and outcome

Table 1: Characteristics of the studies included in the systematic review

Study	Country	Sample size	% males	Mean age, yr	Length of follow-up, mo*	Intervention	Comparison
Chen et al., 2011 ³⁰	China	60	53	38.7	15 (10–20)	TEN	Sling
COTS ¹⁰	Canada	132	78	33.5	12‡	Open reduction plate fixation§	Sling
Figueiredo et al., 2008 ³¹	Brazil	50	78	30.2	24	DCP AI plate	Sling
Judd et al., 2009 ³²	United States	57	91	26.5	12	Modified Hagie pin	Sling
Koch et al., 2008 ³³	Germany	68	66	35.4	19.1 (8–26)	Intramedullary pin	Figure-of-8 bandage
Mirzatooei et al., 2011 ³⁴ †	Iran	60	82	35.6	12	Reconstruction plate on superior surface	Sling
Robinson et al., 2013 ⁵	United Kingdom	200	88	32.4	12	Locking plate	Collar and cuff
Smekal et al., 2009 ³⁵	Austria	68	87	37.7	24	TEN	Sling
Virtanen et al., 2012 ³⁶	Finland	60	87	36.7	12	Reconstruction plate on anterior surface	Sling
Assobhi et al., 2011 ³⁷	Egypt	38	87	31.5	12	AI reconstruction plate	RTEN
Bi et al., 2008 ³⁸	China	201	72	39.8	10.6 (4–21)	Retrograde percutaneous pin	Kirshner pin
Ferran et al., 2010 ³⁹	United Kingdom	133	84	29.2	12	LCDCP	Rockwood pin
Jiang et al., 2012 ⁴⁰ †	China	64	63	42.5	24	LCP	MIPPO
Shen et al., 2008 ⁴¹	China	32	56	44.2	12	Superior reconstruction plate	3-dimensional contoured cortical plate
Lubbert et al., 2008 ⁴²	Netherlands	120	84	NR	12	LIPUS	Placebo

Note: AI = antero-inferior surface, DCP AI = dynamic compression plate in antero-inferior position, LCDCP = limited contact dynamic compression plate, LCP = locking compression plate, LIPUS = low-intensity pulsed ultrasound, MIPPO = minimally invasive percutaneous plate osteosynthesis, NR = not reported, RTEN = retrograde titanium elastic nail, TEN = titanium elastic nail.

*Longest follow-up; reported as mean (range) in 3 studies.^{30,33,38}

†Only enrolled patients with comminuted fractures (3 or more fragments).

‡Data for 2 years in a subsequent publication (Schemitsch et al.⁴³).

§Open reduction and plate fixation (44 patients with limited contact dynamic compression plates; 15 with 3.5-mm reconstruction plates; 4 with precontoured plates, and 4 with other plates).

assessors was described in only 1 study,⁴² it was unclear in 2,^{38,41} and it was not done in the remaining 12 studies.^{5,10,30-37,39,40} Eight studies were considered to be at low risk for attrition bias,^{5,30,33,35-37,40,42} 6 were classified as high risk,^{10,31,32,34,39,41} and 1 was judged as unclear.³⁸ Reporting bias was deemed high in 7 studies,^{10,30,34,35,37,38,41} while 6 were unclear^{31-33,39,40,42} and 2 were considered as low.^{5,36} Agreement between reviewers in the assessment of methodologic quality was excellent (weighted κ value = 0.85). The funnel plot did not suggest publication bias (Figure 3). However, the sample of only 8 studies^{5,10,30-32,34-36} limits interpretability.¹³

Secondary operation

Nonoperative treatment did not confer a greater risk of secondary operation across 8 trials involving 685 patients (RR 1.16, 95% CI 0.58 to 2.35; p for heterogeneity = 0.08, $I^2 = 50%$) (Figure 4). Subgroup analyses suggested an interaction between the type of operative implants (plate versus intramedullary fixation)

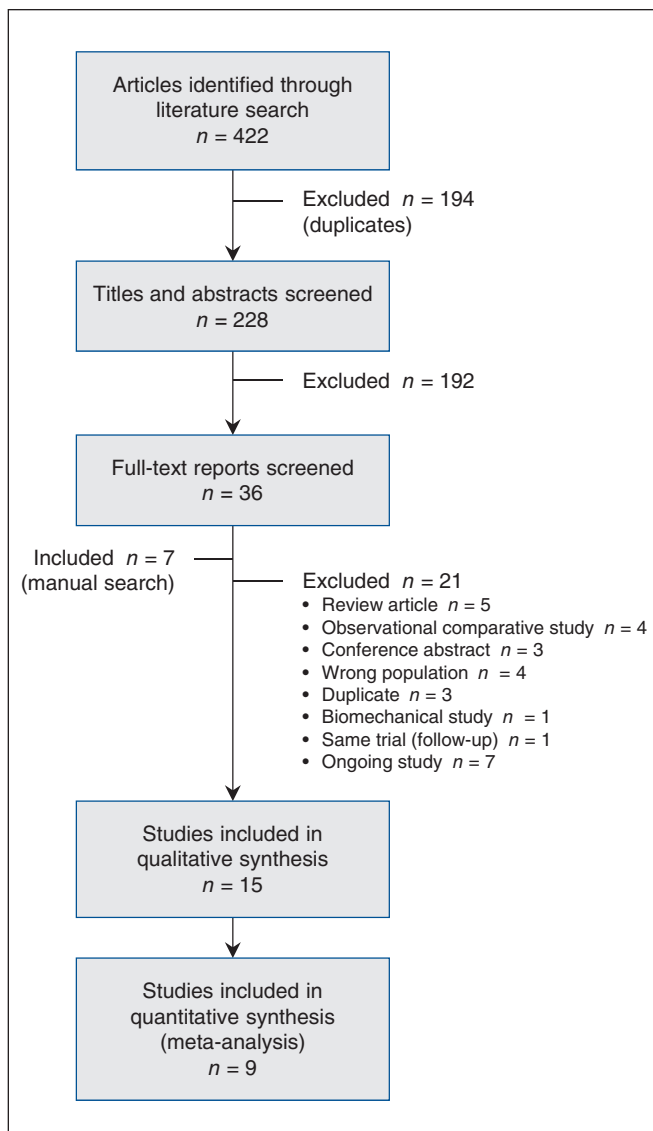


Figure 1: Selection of studies for the meta-analysis.

and the need for secondary operation ($p = 0.05$). These findings were robust to sensitivity testing for trials with missing data (complete case analysis: RR 1.07, 95% CI 0.48 to 2.36; p for heterogeneity = 0.03, $I^2 = 60%$; arm-level assumption analysis: RR 1.12, 95% CI 0.52 to 2.41; p for heterogeneity = 0.02, $I^2 = 62%$).

Reasons for reoperation in the operative group commonly included hardware irritation (54.8%), infection (19%) and implant failure or refracture (19%). Common indications for secondary procedures in nonoperatively managed patients were symptomatic nonunion (57.1%) and symptomatic malunion (28.6%).

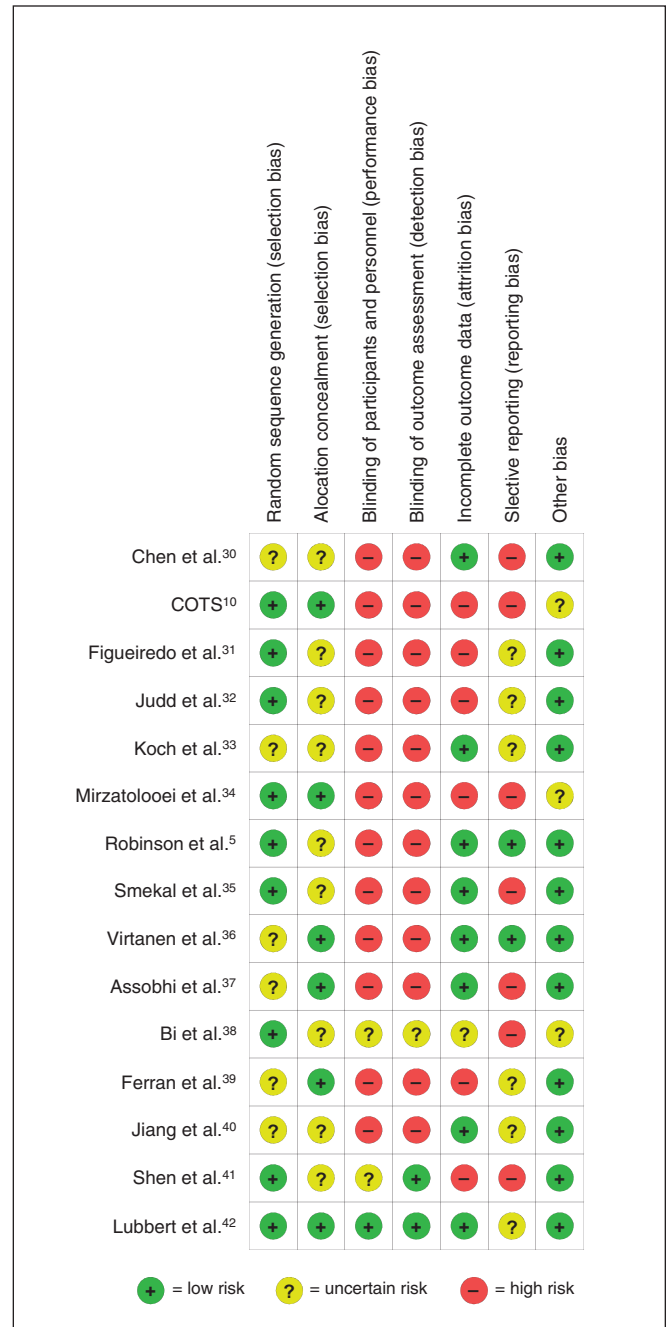


Figure 2: Risk-of-bias assessment of randomized controlled trials included in the meta-analysis. COTS = Canadian Orthopaedic Trauma Society.

All complications

Across 8 studies, there were 77 (23%) complications in 340 patients in the operative group and 88 (26%) complications in 345 patients managed nonoperatively (Table 2). Operative and nonoperative treatments did not differ in complication risk (RR 0.90, 95% CI 0.55 to 1.50; p for heterogeneity = 0.01, $I^2 = 63%$) (Figure 5). Between-trial heterogeneity was not explained by subgroup analysis for type of operative implant ($p = 0.2$). Sensitivity

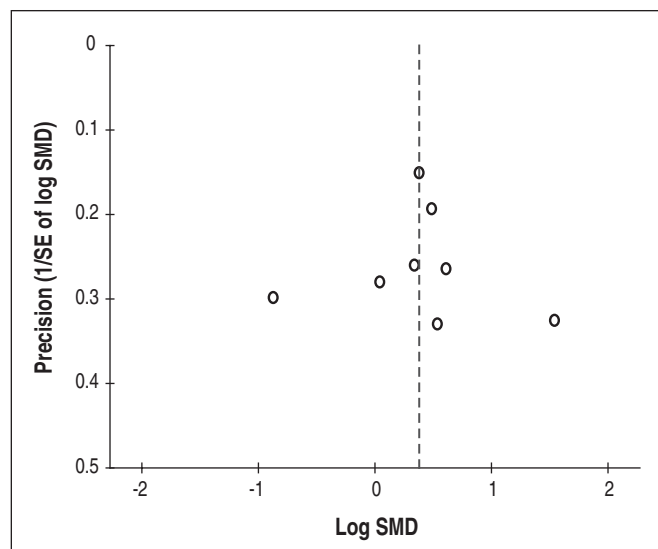


Figure 3: Funnel plot, to assess whether there is evidence of publication bias among trials of operative versus nonoperative treatment of acute displaced midshaft clavicle fractures. The standard normal deviate of the SMD (natural logarithm of the SMD divided by its standard error [SE]) is plotted against the estimate's precision (inverse of the SE). SMD = standardized mean difference.

testing for trials with missing data conferred a similar result (complete case analysis: RR 0.81, 95% CI 0.47 to 1.40; p for heterogeneity = 0.003, $I^2 = 70%$; arm-level assumption analysis: RR 0.86, 95% CI 0.47 to 1.56; p for heterogeneity < 0.001, $I^2 = 78%$).

Functional scores

Seven of the 8 studies included in the pooled analysis evaluated shoulder function at 1 year; the other trial³⁵ assessed it at 2 years. Long-term function favoured operative treatment (SMD 0.38, 95% CI 0.00 to 0.75; p for heterogeneity < 0.001, $I^2 = 79%$). This is equivalent to an estimated DASH score mean difference of 3.5 (95% CI 0.00 to 6.85). The treatment effect failed to exceed the threshold of patient importance based on the MID (10.2 points) (Figure 6). Subgroup analysis to assess the potential risk of selection bias for overall function at 1 year or longer did not differ appreciably from the primary analysis (low risk: $p = 0.06$, $I^2 = 87%$; unclear risk: $p = 0.4$, $I^2 = 73%$). There was residual heterogeneity when we compared high ($p < 0.001$, $I^2 = 0%$) and low ($p = 0.5$, $I^2 = 93%$) risk of attrition bias.

Comparison of operative interventions

Comparison of surgical implants with respect to indications for reoperation, all complications and long-term function have been summarized in Table 3. The functional outcome at 1 year was similar between groups in all trials, showing no significant difference irrespective of the implant used for internal fixation.

Comparison of nonoperative treatments

The available evidence for conservative treatment of acute displaced midshaft clavicle fractures from a placebo-controlled trial of high methodologic quality found no difference in clinical fracture healing between patients receiving low-intensity

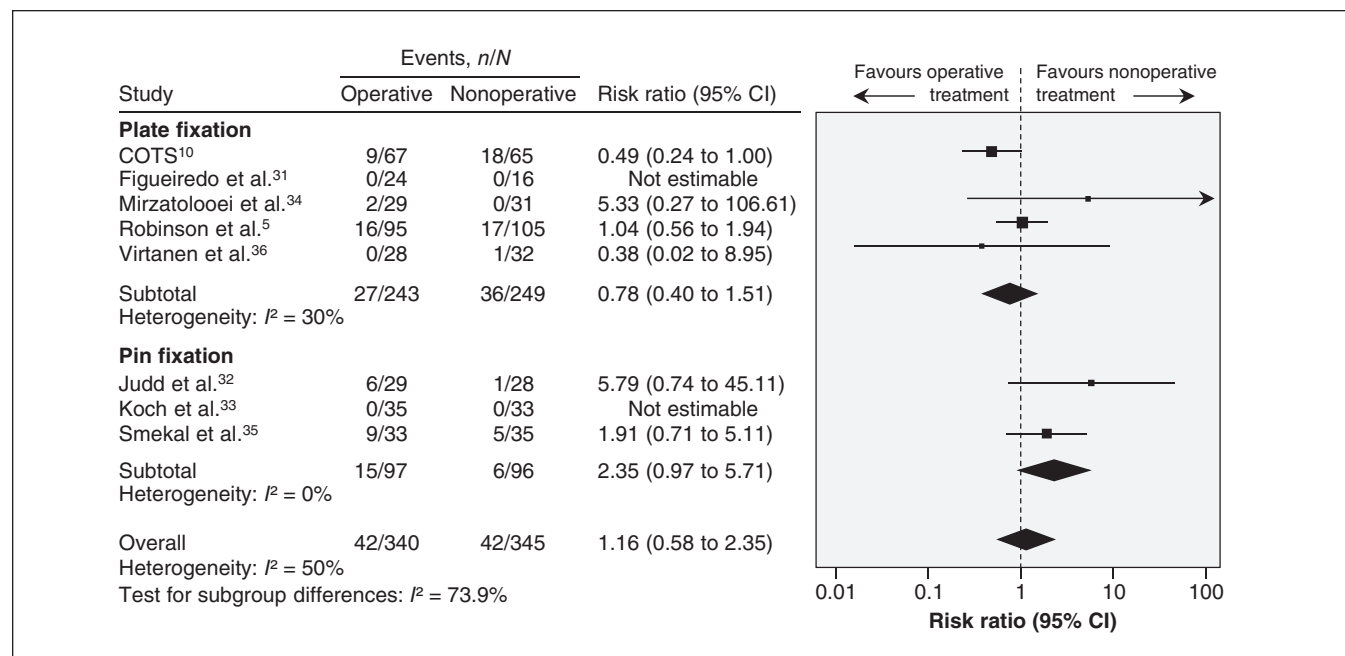


Figure 4: Pooled estimates of secondary operation between operative and nonoperative groups. Values less than 1.0 favour operative treatment. Note: The N values in the study by Figueiredo et al. (24 operative, 16 nonoperative) are the numbers who completed the study and not the numbers initially randomized. CI = confidence interval.

pulsed ultrasound and those given placebo.⁴² Of the 101 patients who completed the trial, 9 (8.9%; 4 placebo, 5 active) underwent subsequent operative treatment with open reduction and internal fixation for fractures that did not heal according to the patients.

Interpretation

The results of our meta-analysis of the relative effects of operative versus nonoperative intervention for acute displaced midshaft clavicle fractures suggest that the incidence of secondary operations and all complications did not differ between the operative and nonoperative groups. There was modest functional improvement at 1 year in the operative group; however, this difference was clinically unimportant. Based on the GRADE criteria (Table 4), the systematic

review and meta-analysis found a lack of high-quality evidence to inform the management of acute displaced midshaft clavicle fractures.

A previous systematic review¹ captured secondary procedures and reported a pooled estimate of effect favouring the operative group (RR 0.38, 95% CI 0.15 to 0.99), a finding that is inconsistent with our review. Our review added a recent RCT and increased the pooled sample size by more than one-third, which likely explains in part the inconsistency. This discrepancy may be further explained by the fact that our review captured hardware irritation and infection as indications for nonroutine secondary procedures, whereas the previous systematic review did not.¹ Although the risk of a secondary procedure was similar between both treatment groups, the reasons for delayed intervention were quite different. Hardware removal because of hardware irritation was the most

Table 2: All complications reported in trials comparing operative and nonoperative interventions

Study	Operative group <i>n</i> = 340		Nonoperative group <i>n</i> = 345	
	Complication*	No. of patients	Complication*	No. of patients
COTS ¹⁰	Operative procedure	9	Operative procedure	18
	Symptomatic nonunion	2	Neurologic symptoms	7
	Neurologic symptoms	8	Complex regional pain syndrome	1
	Abnormality of AC or SC joint	2	Abnormality of AC or SC joint	3
	Other†	2	Other†	2
Figueiredo et al. ³¹	Symptomatic nonunion	2	Symptomatic nonunion	1
	Implant failure	1	Adhesive capsulitis	2
Judd et al. ³²	Operative procedure	6	Operative procedure	1
	Refracture	1	Refracture	1
	Wound infection	3		
	Neurologic symptoms	1		
Koch et al. ³³	NR		NR	
Mirzatooleei et al. ³⁴	Operative procedure	2	Symptomatic malunion	19
	Symptomatic malunion	4	Neurologic symptoms	2
	Early mechanical failure	1		
Robinson et al. ⁵	Operative procedure	16	Operative procedure	17
	Wound infection	2	Rotator cuff impingement	1
	Wound dehiscence	1		
	Rotator cuff impingement	2		
Smekal et al. ³⁵	Operative procedure	9	Operative procedure	5
			Neurologic symptoms	3
Virtanen et al. ³⁶	Refracture	1	Operative procedure	1
	Early mechanical failure	1	Symptomatic malunion	2
	Hardware irritation	1	Refracture	2
Total		77		88

Note: AC = acromioclavicular, NR = not reported, SC = sternoclavicular.
 *Operative procedure = complication that was severe and consequently required secondary operation.
 †Not described.

common indication for a secondary procedure in the operative group, as compared with symptomatic nonunion in the nonoperative group. The latter indication would typically require open reduction and internal fixation with or without bone graft, which may be associated with a higher risk of complications and a longer rehabilitation period than hardware removal because of hardware irritation.

Limitations

Although our population was homogenous in terms of major demographic characteristics, heterogeneity was identified across our key outcomes ($I^2 = 50\%–79\%$). There was substantial heterogeneity in terms of all complications and long-term function, which could be attributed in part to the inclusion of 2 studies^{32,34} in our pooled analysis. Judd and

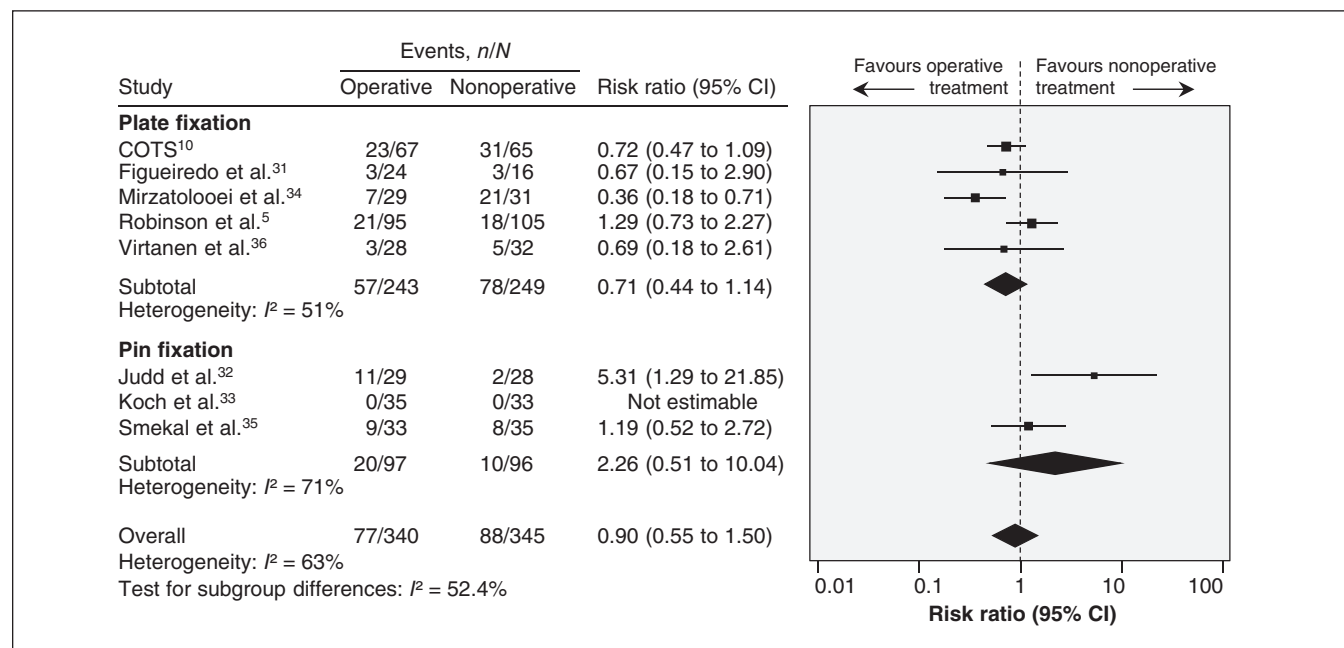


Figure 5: Pooled estimates of all complications between operative and nonoperative groups. Values less than 1.0 favour operative treatment. Note: The N values in the study by Figueiredo et al. (24 operative, 16 nonoperative) are the numbers who completed the study and not the numbers initially randomized. CI = confidence interval.

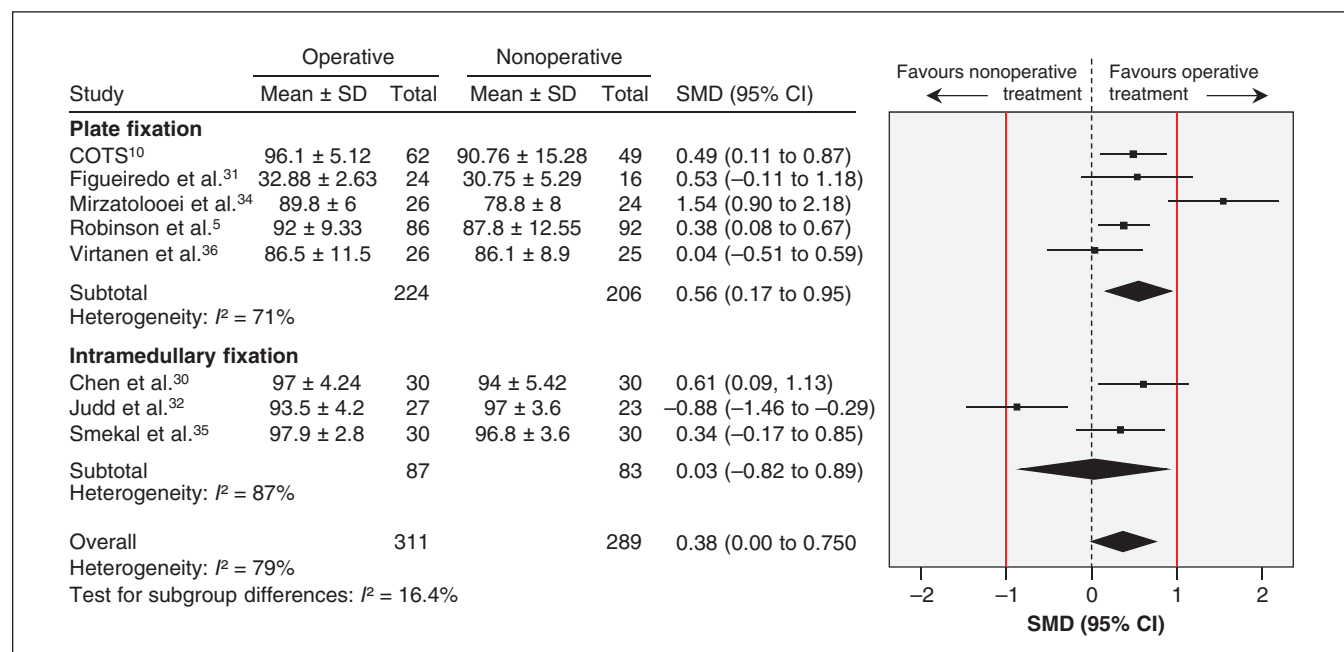


Figure 6: Pooled long-term function (≥ 1 yr) following operative and nonoperative treatment. Standardized mean differences (SMDs) greater than zero favour operative treatment. Red lines show a zone of clinical equivalence based on a minimal important difference of 10.2 points on the Disabilities of the Arm, Shoulder and Hand (DASH) questionnaire. CI = confidence interval, SD = standard deviation.

colleagues³² reported a relatively high complication rate (41%) compared with the other included studies, a difference likely due to the choice of implant. Mirzatooei and colleagues³⁴ included comminution (defined as more than 2 fragments in the fracture site) as an inclusion criterion when assessing patients for eligibility; 33% of the participants in their trial had a fracture with more than 3 fragments. Furthermore, all of the pooled studies excluded open fractures except for this same study, in which 20% of patients had an open clavicle fracture.³⁴ The SMD can be vulnerable to widely varying SDs,¹³ which also may have contributed to the substantial between-study heterogeneity in the pooled analysis for long-term function.

Operative procedures using plates and screws are technically distinct from those using intramedullary devices; however, pooling studies of these techniques separately did not explain the high heterogeneity seen in the primary analyses. Pooling trials at high or unclear risk of selection or detection bias independently from those at low risk did not lower the heterogeneity present in the primary analysis for long-term function. We were unable to perform additional a priori sub-

group analyses, because included studies did not provide sufficient information to explore the effects of time to intervention or comminuted fractures on relevant outcomes.

Seven of the 15 trials included in our review inadequately addressed patients lost to follow-up. Markedly, among the trials comparing operative with nonoperative treatments, a greater number of patients lost to follow-up were in the nonoperative group, which may limit the precision of our estimates of treatment effects and thus overall generalizability.

Implications for practice

Adopting a policy of routine internal fixation for acute displaced midshaft clavicle fractures is contentious, because surgery carries the burden of increased hospital expenditures and inherent surgical complications, including deep or superficial wound infection, hardware irritation, hardware failure or migration, and poor cosmesis of a surgical scar.^{5,9} In a recent retrospective population-based study in Ontario involving 1350 patients who underwent open reduction and internal fixation for a closed isolated midshaft clavicle frac-

Table 3: Summary of secondary operation and complication rates, and functional outcome in trials comparing operative interventions

Study	Secondary operations	Risk ratio (95% CI)	Complications not requiring surgical intervention	Risk ratio (95% CI)	Functional outcome: constant score (1 yr)	Mean difference (95% CI)
Assobhi et al.³⁷						
AI reconstruction plate (<i>n</i> = 19)	1 nonunion 1 wound infection and implant loosening	0.67 (0.13 to 3.55)	1 nonunion	1.00 (0.23 to 4.34)	89.8 (11.3)	-5.60 (-11.21 to 0.01)
RTEN (<i>n</i> = 19)	3 prominent nails		NR		95.5 (5.3)	
Bi et al.³⁸						
Retrograde percutaneous pin (<i>n</i> = 101)	NR	NA	NR	0.11 (0.01 to 2.02)	NR	NA
Kirshner pin (<i>n</i> = 100)	NR		4 nonunions		NR	
Ferran et al.³⁹						
Rockwood pin (<i>n</i> = 17)	1 implant loosening	0.22 (0.06 to 0.88)	NR	0.22 (0.06 to 0.88)	92.1 (6)*	3.4 (-2.02 to 8.82)
LCDCP (<i>n</i> = 15)	3 superficial infections 1 persistent pain 4 hardware irritation		NR		88.7 (9.1)*	
Jiang et al.⁴⁰						
MIPPO (<i>n</i> = 32)	NR	NA	NR	NA	96†	0.30
LCP (<i>n</i> = 32)	NR		NR		95.7†	(-4.70 to 5.30)
Shen et al.⁴¹						
3D contoured cortical plate (<i>n</i> = 67)	1 delayed union	0.12 (0.02 to 0.96)	3 "symptomatic patients"	0.20 (0.06 to 0.65)	NR	NA
Superior reconstruction plate (<i>n</i> = 66)	8 delayed unions		15 "symptomatic patients"			

Note: AI = antero-inferior surface, LCDCP = limited contact dynamic compression plate, LCP = locking compression plate, MIPPO = minimally invasive percutaneous plate osteosynthesis, NA = not applicable, NR = not reported, RTEN = retrograde titanium elastic nail.
*Unclear as to whether this was at 1-year assessment.
†No standard deviation reported; means were abstracted from graphical analyses.

ture,⁴⁴ the reoperation rate was 24.6%, which is about twice as high as our finding. Fifty percent of the patients had their hardware removed after 12 months (median 12 mo, interquartile range 5.8 to 16.1 mo), whereas more than half of the trials included in our meta-analysis had a follow-up period of only 12 months.

There are important differences in design characteristics between observational studies and randomized trials that may be responsible for contradictory estimates of treatment effects. First, infrequent events and long-term clinical outcomes are often difficult to study in randomized trials and may be more suitably investigated in large observational studies.⁴⁵ Second, it is plausible that surgeons involved in most trials may have substantial generic surgical expertise and expertise with the intervention under study, which may not represent the skill level of the surgical community in which the intervention will be implemented.⁴⁶ Despite the obvious discrepancy between the observational studies and RCTs in terms of reoperation, it is incumbent upon us to recognize the complementary roles of both sources of information and understand that the complete body of evidence could have profound clinical implications.

Implications for research

Recurrent study design limitations, including small samples, lack of blinding and loss to follow-up, must be overcome to improve the quality of evidence from future RCTs. Unified evaluation criteria for outcomes such as nonunion and malunion should be applied to all trials evaluating interventions to treat these fractures. The most recent RCT included in

our review evaluated absence from work and found that, although the timing of return to work was dependent on the nature of the patient's work, no significant differences were found between the study groups in terms of total time off work following injury ($p = 0.7$).⁵ Therefore, if long-term function is seemingly similar between treatment groups, further investigation should aim to determine whether early functional improvements (< 6 mo) differ significantly between operative and nonoperative groups.

The trials included in our review did not provide sufficient evidence to suggest which patients might benefit the most from surgical treatment. It remains unclear whether certain fracture characteristics such as shortening, displacement or comminution can reliably predict patient-focused functional outcomes.⁴⁷ A study assessing the reliability of fellowship-trained shoulder and sports medicine orthopedic surgeons in classifying midshaft clavicle fractures via standard plain radiographs and the agreement in treating the fractures showed moderate to strong agreement for degree of displacement and for comminution; however, the standard plain radiographs were insufficient to reliably determine the degree of shortening of clavicle fractures and the need for surgery among the surgeons.⁴⁸ Further investigations are required to develop better criteria to avoid under- or overestimating fracture severity. Future trials should aim to better identify the subgroup of patients who might benefit from primary surgical intervention and to establish optimal surgical indications. Their findings would help focus the use of surgical resources on appropriate candidates and prevent undertreatment of the injury nonoperatively.

Table 4: GRADE summary of findings for operative compared with nonoperative treatment for acute displaced midshaft clavicle fractures

Patient or population: Patients with an acute displaced midshaft clavicle fracture Intervention: Operative treatment (plate or intramedullary device) Comparison: Nonoperative treatment (standard sling, figure of 8 dressing, or a collar and cuff) Outcomes: Secondary operations, all other complications, long-term function			
Outcome	No. of participants (studies)	Anticipated effects, risk with operative treatment (95% CI)	GRADE quality of evidence
Secondary operation Follow-up: 12 mo	685 (8)	Evidence suggested higher incidence of secondary operation in operative group (RR 1.16, 95% CI 0.58 to 2.35), but difference was not statistically significant	Low*‡
All other complications Follow-up: 12 mo	685 (8)	Evidence suggested slightly lower number of complications in the operative group (RR 0.9, 95% CI 0.55 to 1.5), but difference was not statistically significant	Low*‡
Long-term function Follow-up: (≥ 1 yr)	611 (8)	Mean long-term shoulder function was 0.38 SDs higher in operative group (0.22 lower to 0.54 higher)†	Very low*§
<p>Note: CI = confidence interval, GRADE = Grading of Recommendations Assessment, Development, and Evaluation, MID = minimal important difference, RR = risk ratio, SD = standard deviation. *Downgraded because of risk of bias (lack of blinding study personnel, unclear reporting of allocation concealment and sequence generation). †Effect failed to exceed MID (smallest effect that an informed patient would perceive as beneficial enough to justify a change in management). ‡Downgraded because of fragility of few events. §Downgraded for imprecision and inconsistency.</p>			
<p>GRADE Working Group grades of evidence: High quality: Further research is very unlikely to change our confidence in the estimate of effect. Moderate quality: Further research is likely to have an important impact on our confidence in the estimate of effect and may change the estimate. Low quality: Further research is very likely to have an important impact on our confidence in the estimate of effect and is likely to change the estimate. Very low quality: We are very uncertain about the estimate.</p>			

Conclusion

Current evidence does not support the routine use of internal fixation for the treatment of displaced midshaft clavicle fractures. Evidence for the type of implant or approach to nonoperative treatment was inconclusive, and complication rates were high regardless of the management approach.

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