

Site of hospital readmission and mortality: a population-based retrospective cohort study

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Abstract

Background: Unplanned hospital readmission is a complex process, particularly if the patient is readmitted to an acute care institution other than the original hospital. This study tested the hypothesis that readmission to an alternative hospital is associated with increased mortality compared with readmission to the original hospital.

Methods: We performed a population-based retrospective cohort analysis set between 1995 and 2010 for all 21 acute care adult general hospitals in the Greater Toronto and Hamilton Area. Participants were consecutive adults (age ≥ 18 yr) readmitted through the emergency department within 30 days after hospital discharge. The primary outcome measure was all-cause mortality within 30 days after readmission.

Results: Of the 198 149 patients included in the study, 38 134 (19.2%) died within 30 days after readmission. Patients readmitted to an alternative hospital were more likely than those readmitted to the original hospital to be older, reside in a chronic-care facility and arrive by ambulance. Alternative-hospital readmission was associated with a higher risk of death within 30 days (22.3% v. 18.6%, p < 0.001; odds ratio [OR] 1.26, 95% confidence interval [CI] 1.23–1.30). The increased risk was substantially less after adjustment for patient- and hospital-level covariables (adjusted OR 1.06, 95% CI 1.02–1.10). Unadjusted Kaplan–Meier survival curves separated early and the absolute difference in mortality continued throughout the entire 1-year follow-up period, but no difference between groups was observed based on adjusted survival analyses.

Interpretation: Among patients readmitted within 30 days after discharge, readmission to an alternative hospital was associated with a higher risk of death than readmission to the original hospital. Whether this adverse prognosis reflects a true causal relation or residual confounding is unknown.

ach year, about 4.5 million people in the United States and about 0.2 million Canadians undergo unplanned readmission to hospital within 30 days after hospital discharge. ^{1,2} Unplanned readmissions are costly, challenging to predict, difficult to avoid and associated with increased mortality. ³⁻⁹ Although most readmitted patients return to the original hospital, about a fifth spend their initial ("primary") and subsequent ("secondary") stays in different hospitals. ¹⁰ Clinical experience suggests that unfamiliar patients can pose challenges, but 1 study found no large difference in mortality between patients readmitted to original and alternative hospitals. ¹¹

Patients readmitted to an alternative hospital might be disadvantaged by limited access to their primary admission medical records, delayed treatment, discontinuity of care and increased exposure to nosocomial pathogens. ¹²⁻¹⁸ On the other hand, clinicians in an alternative hospital may be less likely to replicate errors committed during the primary admission, less influenced by previously affixed diagnostic labels and less

dispirited by a patient's return to hospital.¹⁹ In addition, procedures and protocols at the alternative hospital may intercept problems arising in the original hospital and facilitate rectification of prior medical errors.²⁰ These multiple countervailing forces have an uncertain aggregate influence on patient outcomes.

Whether a patient is readmitted to the original hospital or an alternative is determined by factors that include hospital proximity, specialist availability, ambulance referral patterns, local traffic conditions, anticipated emergency department delays, institutional reputation and patient preference.^{21–24}

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The complexity and heterogeneity of these factors make a prospective randomized trial logistically challenging and ethically dubious. Thus, we conducted a population-based retrospective cohort analysis using linked administrative databases to test whether readmission to an alternative hospital is associated with a higher risk of death than readmission to the original hospital.

Methods

Setting

We focused on a large metropolitan area, reasoning that readmissions to alternative hospitals in rural regions would be uncommon, subject to practicalities, such as distances, and confounded by referrals for urban subspecialty care. We selected Ontario's Greater Toronto and Hamilton Area because it is Canada's largest contiguous urban region. ²⁵ At the midpoint of our study (2002), the area covered an area of 8241 km² and had a population of 5.6 million. ²⁶ Throughout the study interval, Ontario residents had universal health insurance that provided access to medical care that was widely available, publicly funded and free at the point of service. ²⁷ The health care system within the Greater Toronto and Hamilton Area included a full array of primary, secondary, tertiary and quaternary care, and no major changes to hospital financing occurred during the study.

Hospitals

We performed an individual-level analysis of patients whose primary and secondary admissions to hospital both occurred within the study area. This approach minimized confounding by excluding patients referred from rural regions for subspecialty care. We first identified all acute care hospitals operating in the study area between 1995 and 2010. Individual inpatient acute care sites sharing the same facility identification number were classified as a multisite hospital because their shared data information systems, personnel, protocols and governance made that the smallest meaningful functional unit for analysis. Individual inpatient acute care sites with no sister sites were classified as single-site hospitals. We excluded hospitals without adult inpatient acute care beds and hospitals restricted to pediatric, elective surgical, rehabilitative, respite or palliative care. Inpatient acute care sites that joined a multisite hospital during the study period were excluded before the merger date.

Patients

We identified adult patients aged 18 years or older who were associated with 1 or more eligible hospital readmissions during the study interval. A readmission was eligible if the patient was readmitted to hospital through the emergency department between Jan. 1, 1995, and Dec. 31, 2010; the time between initial hospital discharge and subsequent hospital readmission was 1–30 days; and both the primary and secondary admissions occurred in qualifying hospitals in the study area. A readmission was ineligible if the most responsible diagnosis at the primary or secondary admission was

related to pregnancy (International Classification of Diseases, 9th Revision [ICD-9] codes 630–676; International Statistical Classification of Diseases and Related Health Problems, Tenth Revision, Canada [ICD-10-CA] codes O00–O99) or a major psychiatric disorder (ICD-9 codes 295–300; ICD-10-CA codes F20–F48). Same-day readmissions were excluded to avoid misidentifying hospital-to-hospital transfers as readmissions. For patients with multiple readmissions, only the most recent eligible readmission was analyzed, so that each individual contributed only once to analyses. (See eFigure 1 in Appendix 1 available at www.cmajopen.ca/content/2/2/E77/suppl/DC1 for a schematic representation of the study.)

The Research Ethics Board of the Sunnybrook Research Institute approved the study and waived the requirement for individual consent.

Variables

The primary outcome was all-cause mortality within 30 days after hospital readmission, as ascertained from the official government vital statistics record for Ontario. Secondary outcomes were in-hospital case fatality and all-cause mortality by 90, 180 and 365 days after readmission. The predictor of interest was the site of readmission, categorized as original hospital or alternative hospital. Original-hospital readmissions were defined as those where both primary and secondary admissions occurred at the same hospital. Alternative-hospital readmissions were those where primary and secondary admissions occurred at different hospitals.

Covariables used for adjusted analyses were selected based on earlier research and included patient demographics, neighbourhood average household income (divided into quintiles), rural residence (yes, no), year of readmission, Charlson Comorbidity Index (integer), primary length of stay (days), most responsible diagnosis for secondary hospital admission, arrival by ambulance at the time of readmission (yes, no), residence in a chronic care facility at time of readmission (yes, no), total number of admissions to hospital in the year before readmission (integer) and total number of physician clinic visits in the year before readmission (integer).5,28 We included the hospital-free interval, defined as the number of days between hospital discharge and readmission. On the basis of earlier reports, we also examined annual case volume (integer) and sector (designated as academic sector or community sector by system-level hospital reports) of the readmission hospital.²⁹⁻³³ Variables were modelled as continuous values unless otherwise indicated but are presented in categories to facilitate interpretation.

In the Canadian setting, the condition most responsible for a patient's hospital stay is reported as the most responsible diagnosis.³⁴ In our study, the most responsible diagnosis for each admission was coded according to the ICD-9 before Apr. 1, 2002, and the ICD-10-CA after Apr. 1, 2002. We adapted the 285 single-level diagnostic categories in the Agency for Healthcare Research and Quality's Clinical Classifications Software to cluster individual diagnostic codes into 30 mutually exclusive, clinically meaningful diagnostic categories (eTable 1 in Appendix 1).³⁵

Data sources

Encrypted patient-level data were obtained from populationbased, linked administrative databases that have been used extensively in research^{17,36,37} and validated.³⁸⁻⁴¹ Information on outpatient clinic visits was obtained from the Ontario Health Insurance Plan database. Information on hospital admissions was obtained from the Canadian Institute for Health Information database. Demographics, vital status and date of death were obtained from official government records. Socioeconomic status was estimated using average household income at the census tract level based on the 2006 Canadian census. Hospital structural and operational data (including designation as an academic or community hospital) were obtained from reports published at or near the study midpoint.^{29,42–44} The study interval extended from 1995 to 2010 to include the most recent year for which data were available and to reflect an interval with no major changes in hospital financing.

Missing data for neighbourhood average household income and rural residence were coded explicitly as missing and retained for analysis (less than 0.1% of patients). Patient-level data were otherwise complete.

Statistical analysis

The primary analysis compared 30-day mortality after alternative-hospital readmission with 30-day mortality after original-hospital readmission. All analyses were at the level of the individual patient. We compared unadjusted risks using a χ^2 test and then adjusted for all patient- and hospital-level covariables listed above using a logistic regression model that applied general estimating equations (GEE) with an exchangeable correlation structure to account for clustering of individuals within hospitals.45 In a subsequent patient-level analysis, we stratified patients by secondary hospital, calculated unadjusted odds ratios (ORs) for each secondary hospital stratum, and then calculated adjusted ORs for each secondary hospital stratum by fitting separate logistic regression models to each hospital to account for the individual patient characteristics listed above. In further analyses, we stratified the pooled cohort by patient characteristics and used unadjusted univariable GEE to explore potential effect modification.

Kaplan-Meier survival curves were generated for unadjusted survival up to 1 year and tested using the log-rank test. Survival analysis was done using a Cox proportional-hazards analysis adjusted for all patient- and hospital-level characteristics listed above. In-hospital case fatality and death by 90, 180 and 365 days were examined using the multivariable GEE logistic regression model developed for the primary outcome. One sensitivity analysis retained only the first (rather than most recent) eligible readmission for each patient. A second sensitivity analysis examined the influence of analyzing according to primary (rather than secondary) hospital. A third sensitivity analysis defined all patients readmitted within 48 hours as hospital-tohospital transfers and excluded those patients from analysis. A post hoc exploratory analysis examined the relative influence of distinct hospital-sector transitions (Appendix 1). We performed all analyses using 2-sided statistical tests at the 5% level of significance using SAS, version 9.2 (SAS Institute Inc., Cary, NC).

Results

Hospital and patient characteristics

A total of 43 inpatient acute care sites grouped within 28 distinct hospitals were active in the Greater Toronto and Hamilton Area during the study period. Seven single-site hospitals provided highly restricted subspecialty care and were excluded from analysis (3 provided only elective surgical services, 3 provided only palliative care and 1 provided only pediatric care), leaving 21 qualifying acute care hospitals (eFigure 2 in Appendix 1). Hospital characteristics at the midpoint of the study can be found in eTables 2 and 3 in Appendix 1.

We identified 2 448 759 patients who were admitted at least once to a qualifying acute care hospital in the study area during the 16-year study period; of these, 198 228 (8.1%) subsequently had an eligible readmission to a hospital in the study area (Figure 1). We excluded 79 patients, whose date of death was erroneously recorded as occurring before the readmission date, resulting in a cohort of 198 149 patients. Of these, 58 460 (29.5%) patients had more than 1 eligible readmission (eFigure 3 in Appendix 1), but only the most recent readmission episode for each patient was retained in the primary analysis. Less than 0.003% of data were missing for age, sex, neighbourhood average household income, prior physician visits, prior admissions to hospital and most responsible diagnosis at primary and secondary admission.

Of the 198 149 patients in our cohort, 161 974 (81.7%) were readmitted to the original hospital and 36 175 (18.3%) to an alternative hospital. Compared with patients readmitted to the original hospital, those readmitted to an alternative hospital were more likely to be older, male, burdened by more comorbidity (as measured by hospital admissions and clinic visits in the previous year and by Charlson Comorbidity Index), a resident of a chronic care facility and brought to hospital by

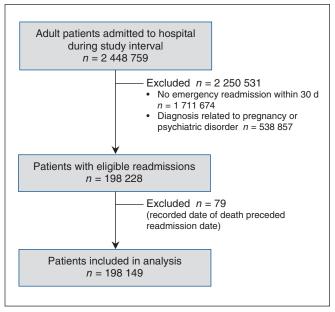


Figure 1: Selection of patients readmitted to hospital in the Greater Toronto and Hamilton Area within 30 days after discharge.

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ambulance (Table 1). Compared with original-hospital readmissions, alternative-hospital readmissions were somewhat more likely to occur at community hospitals or at those with a lower annual case volume (eTable 6 in Appendix 1).

Short-term mortality

Compared with patients readmitted to the original hospital, those readmitted to an alternative hospital were significantly more likely to die within 30 days after readmission (8072 [22.3%] of 36 175 patients v. 30 062 [18.6%] of 161 974 patients; χ^2 test, p < 0.001). This difference is equivalent to an

unadjusted OR of 1.26 (95% confidence interval [CI] 1.23–1.30). Adjustment for patient and hospital factors substantially attenuated this association (adjusted OR 1.06, 95% CI 1.02–1.10). Additional independent predictors of death appear in Table 2. In our final model, the relative increase in risk of death within 30 days attributable to alternative-hospital readmission was generally smaller than the risk associated with sepsis or pneumonia (Table 2). The within-hospital correlation coefficient from our primary analysis was 0.12.

Alternative-hospital readmission was associated with a significant increase in the adjusted odds of death within 30 days

| | Readmission hospit | | |
|---|----------------------------------|-----------------------------------|---------|
| Characteristic | Original hospital n = 161 974 | Alternative hospital $n = 36 175$ | p value |
| Age, yr* | | | |
| < 65 | 63 578 (39.3) | 13 064 (36.1) | < 0.001 |
| 65–74 | 33 471 (20.7) | 7 602 (21.0) | < 0.001 |
| 75–84 | 41 941 (25.9) | 9 650 (26.7) | < 0.001 |
| ≥ 85 | 22 984 (14.2) | 5 859 (16.2) | < 0.001 |
| Sex, female | 80 941 (50.0) | 17 654 (48.8) | < 0.001 |
| Neighbourhood average household income† | | | |
| Highest quintile | 29 375 (18.1) | 6 675 (18.5) | < 0.001 |
| Next to highest quintile | 29 277 (18.1) | 6 381 (17.6) | < 0.001 |
| Middle quintile | 30 910 (19.1) | 6 609 (18.3) | < 0.001 |
| Next to lowest quintile | 35 262 (21.8) | 7 653 (21.2) | < 0.001 |
| Lowest quintile | 36 696 (22.7) | 8 732 (24.1) | < 0.001 |
| Rural residence† | 3 198 (2.0) | 465 (1.3) | < 0.001 |
| Additional admissions to hospital in previous year* | 81 549 (50.3) | 19 428 (53.7) | < 0.001 |
| ≥ 7 physician clinic visits in previous year* | 125 726 (77.6) | 28 540 (78.9) | < 0.001 |
| Charlson Comorbidity Index ≥ 2* | 84 955 (52.4) | 19 704 (54.5) | < 0.001 |
| Most responsible diagnosis for primary admission‡ | | | |
| Malignant neoplasm | 19 590 (12.1) | 4 303 (11.9) | 0.3 |
| Injury | 10 666 (6.6) | 2 724 (7.5) | < 0.001 |
| Acute coronary syndrome | 9 000 (5.6) | 2 115 (5.8) | 0.030 |
| Heart failure | 9 258 (5.7) | 1 856 (5.1) | < 0.001 |
| Pneumonia | 6 696 (4.1) | 1 588 (4.4) | 0.03 |
| Obstructive lung disease | 6 742 (4.2) | 1 316 (3.6) | < 0.001 |
| Chronic ischemic heart disease | 3 168 (2.0) | 2 440 (6.7) | < 0.001 |
| Cardiac arrhythmia | 4 318 (2.7) | 824 (2.3) | < 0.001 |
| Cerebrovascular disease | 3 708 (2.3) | 1 135 (3.1) | < 0.001 |
| Miscellaneous | 49 237 (30.4) | 10 652 (29.4) | < 0.001 |
| Primary length of stay* greater than cohort median | 92 533 (57.1) | 20 843 (57.6) | 0.09 |
| Hospital-free interval* greater than cohort median | 80 749 (49.9) | 20 321 (56.2) | < 0.001 |
| Residence in chronic care facility at readmission | 16 479 (10.2) | 5 446 (15.1) | < 0.001 |
| Arrival by ambulance at readmission | 70 523 (43.5) | 20 677 (57.2) | < 0.001 |

^{*}Analyzed as continuous data but presented here in categories to facilitate interpretation.

[†]Missing data were coded as such and included in analyses; however, missing data were uncommon (< 0.003% of all data) and are not presented here.

[‡]The 10 most common diagnostic categories are presented here. For the complete list of diagnostic categories and their frequency at primary and secondary admissions to hospital, see eTables 4 and 5 in Appendix 1.



for 4 of 6 academic secondary hospitals. For all other secondary hospitals, there was no significant difference in the adjusted risk of death between original-hospital and alternative-hospital readmissions (Figure 2; eTable 7 in Appendix 1). Post hoc stratification of the multivariable GEE model by sector of the secondary hospital showed a significant association between 30-day mortality and alternative-hospital readmission for academic hospitals (adjusted OR 1.17, 95% CI 1.15–1.19), but not for community hospitals (adjusted OR 1.00, 95% CI 0.97–1.04). Results of an exploratory analysis examining specific hospital sector transitions yielded similar findings (eTable 9 in Appendix 1).

The unadjusted association between alternative-hospital readmission and increased mortality was present among

patients of all age groups, all neighbourhood household income quintiles and both sexes (Table 3). The unadjusted relative increase in risk was accentuated among women and among patients with less comorbidity, fewer admissions to hospital and a residence other than a chronic care facility.

Delayed mortality

Unadjusted survival analysis showed that the risk associated with alternative-hospital readmission was equal to an absolute difference in mortality of about 4% that was sustained throughout the 1-year follow-up period (Figure 3; log rank test p < 0.001). In contrast, adjusted survival analysis found no difference between the groups (hazard ratio 1.01, 95% CI 0.99–1.02). Adjusted analyses at specific times revealed that

| Factor | Adjusted OR* (95% CI) | Factor | Adjusted OR* (95% CI) |
|---|--------------------------|-------------------------------------|--------------------------|
| Age, yr† | | Residence in chronic care facility | |
| < 65 (ref) | 1.00 | at readmission | |
| 65–74 | 1.26 (1.23-1.30) | No (ref) | 1.00 |
| 75–84 | 1.49 (1.43-1.55) | Yes | 1.31 (1.25–1.37) |
| ≥ 85 | 2.05 (1.90-2.21) | Most responsible diagnosis | |
| Sex | | for readmission§ | |
| Female (ref) | 1.00 | Cardiac arrest | 3.00 (5.70-9.18) |
| Male | 1.15 (1.13–1.16) | Malignant neoplasm | 3.57 (3.16-4.03) |
| Neighbourhood household income‡ | - (| Sepsis | 3.10 (2.72–3.53) |
| Highest quintile (ref) | 1.00 | Liver disease | 2.29 (2.03-2.59) |
| Next to highest quintile | 0.99 (0.96–1.03) | Pneumonia | 1.87 (1.69-2.07) |
| Middle quintile | 1.01 (0.98–1.04) | Nephrolithiasis | 0.44 (0.41-0.47) |
| Next to lowest quintile | 1.03 (0.98–1.09) | Nonspecific abdominal pain | 0.40 (0.36-0.43) |
| Lowest quintile | 1.01 (0.98–1.05) | Nonspecific chest pain | 0.26 (0.24-0.28) |
| Admissions to hospital in | (0.00 | Syncope and collapse | 0.24 (0.22-0.27) |
| previous year† | | Prostatic hypertrophy | 0.24 (0.20-0.29) |
| 1 (ref) | 1.00 | Miscellaneous (ref) | 1.00 |
| ≥2 | 1.11 (1.09–1.14) | Arrival by ambulance at readmission | |
| Physician clinic visits in | () | No (ref) | 1.00 |
| previous year† | | Yes | 1.95 (1.88-2.01) |
| ≤ 6 (ref) | 1.00 | Year of readmission† | |
| ≥ 7 | 0.97 (0.94–1.00) | 1994–1999 (ref) | 1.00 |
| Charlson Comorbidity Index† | 0.07 (0.01 1.00) | 2000–2005 | 0.95 (0.92-0.99) |
| ≤ 1 (ref) | 1.00 | 2006–2010 | 0.87 (0.83-0.92) |
| ≥ 1 (161) ≥ 2 | 1.87 (1.80–1.94) | Sector of secondary hospital | |
| | 1.07 (1.00–1.94) | Community (ref) | 1.00 |
| Primary length of stay† longer than cohort median | | Academic | 0.70 (0.59-0.82) |
| No (ref) | 1.00 | Case volume of secondary hospital† | |
| Yes | 1.20 (1.17–1.23) | Lower (ref) | 1.00 |
| Hospital-free interval† longer | (1.17 1.20) | Higher | 1.04 (0.86–1.25) |
| than cohort median | | | |
| No (ref) | 1.00 | | |
| Yes | 1.03 (1.00–1.05) | | |

Note: CI = confidence interval, OR = odds ratio, ref = reference group.

^{*}Adjusted for age, sex, neighbourhood average household income, rural residence, year of readmission, Charlson Comorbidity Index, primary length of stay, most responsible diagnosis at readmission to hospital, arrival by ambulance at readmission, residence in chronic care facility at readmission, admission to hospital in previous year, physician clinic visits in previous year, hospital-free interval, case volume of secondary hospital and sector of secondary hospital.

[†]Analyzed as continuous data but presented here in categories to facilitate interpretation.

‡Missing data were coded as such and included in analyses; however, missing data were uncommon (< 0.003% of all data) and are not presented here.

^{\$}The 5 diagnostic categories with the greatest effect on mortality and the 5 with the least effect are presented here. For the complete list, see eTable 8 in Appendix 1.

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the adjusted odds of death at 90, 180 or 365 days after readmission were no different for original- and alternative-hospital readmissions (eTable 11 in Appendix 1).

Sensitivity analyses

Retaining only the first rather than the most recent readmission for each patient resulted in a similar summary effect estimate for the primary outcome: 20 571 (12.7%) of the 161 955 patients readmitted to the original hospital died, as compared

with 5780 (16.0%) of the 36 194 patients readmitted to an alternative hospital died (unadjusted OR 1.30, 95% CI 1.27–1.35; adjusted OR 1.08, 95% CI 1.04–1.12). Analyzing by primary rather than secondary hospital had little effect on the summary effect measure generated by the multivariable GEE model (adjusted OR 1.08, 95% CI 1.01–1.16), despite more pronounced influences on the unadjusted hospital-specific effect estimates (eFigure 4 in Appendix 1). Extending the definition of hospital-to-hospital transfer to exclude patients

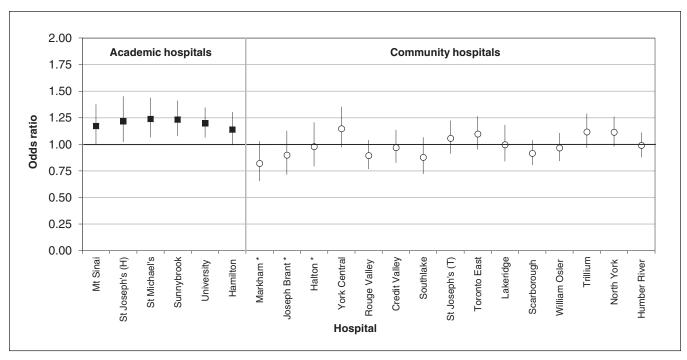


Figure 2: Hospital-specific adjusted odds of death within 30 days after readmission among patients readmitted to an alternative hospital compared with those readmitted to the original hospital. Values greater than 1.0 indicate an increased risk of death. Error bars represent 95% confidence intervals. *Multivariable logistic regression produced unstable estimates due to nonconvergence.

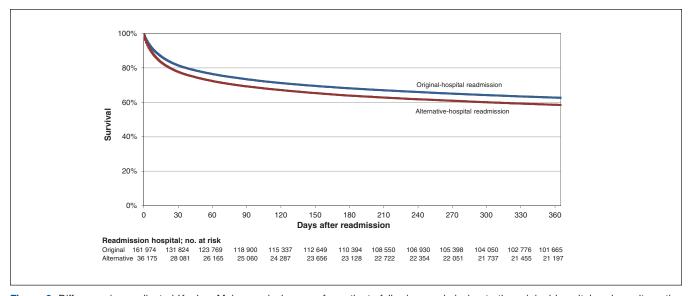


Figure 3: Difference in unadjusted Kaplan–Meier survival curves for patients following readmission to the original hospital and an alternative hospital (p < 0.001).



readmitted within 48 hours after discharge from the primary hospital resulted in a smaller cohort (n = 187 446), but did not influence the primary outcome summary effect measure (adjusted OR 1.05, 95% CI 1.02–1.09).

Interpretation

In this population-based observational cohort study, readmission to an alternative hospital rather than the original hospital was associated with increased patient mortality that was attenuated substantially after adjustment for patient- and hospitallevel covariables. One interpretation of these findings is that alternative-hospital readmission can compromise patient safety, in accord with past studies examining discontinuous care. 46-48 Another interpretation is that increased patient mortality reflects residual confounding by unmeasured covariables that differed for original- and alternative-hospital patients in our cohort. The latter interpretation acknowledges that conditions with relatively poor prognoses might also predispose patients to alternative-hospital readmission and that system factors promoting alternative-hospital readmission (e.g., ambulance diversion or prolonged wait times in the emergency department) potentially delay care and adversely affect patient outcomes.^{49,50} Given the potential importance of our findings in terms of patient outcomes, further examination of the underlying association between alternative-hospital readmission and mortality is warranted.

Our exploratory analyses found an association between alternative-hospital readmission and mortality for secondary hospitals in the academic sector, but not for those in the community sector. Exploratory comparisons of specific hospital sector transitions also suggested that sector influences the association between readmission to an alternative hospital and mortality (eTable 9 in Appendix 1). One explanation for this finding is that academic hospitals care for highly complex patients who might be particularly vulnerable to discontinuities in care.31 A second explanation is the confounding that might be introduced if patients with severe or unusual illnesses were directed to subspecialty academic centres for readmission. A third explanation focuses on differences in organizational culture between academic and community hospitals.⁵⁰ Further research, including organizational network analysis, may help clarify this finding. 51,52

Several features distinguish our study from past research.¹¹ We defined alternative-hospital versus original-hospital readmission in advance as the primary predictor of interest. Our patient cohort was population-based, universally insured, selected over an extended interval and substantial in size. We focused on a large contiguous metropolitan area and excluded rural hospitals with potentially idiosyncratic readmission patterns. Each hospital in our study contributed a large number of patients (average of 9455 per hospital), which allowed analyses to account for actual readmission site. Finally, a high proportion of our cohort died within 30 days after readmission, which suggests that our patients had a substantial burden of comorbidity, a high acuity of presenting illness or some combination of these factors.

Table 3: Unadjusted odds of death within 30 days after readmission to an alternative hospital compared with readmission to the original hospital, stratified by patient characteristics

| Patient characteristic | Unadjusted OR* (95% CI |
|--|--|
| Age, yr | |
| < 65 | 1.36 (1.27–1.46) |
| 65–74 | 1.22 (1.11–1.33) |
| 75–84 | 1.16 (1.09–1.23) |
| ≥ 85 | 1.15 (1.07–1.23) |
| Sex | |
| Female | 1.35 (1.26–1.45) |
| Male | 1.16 (1.08–1.25) |
| Neighbourhood household income | - () |
| Highest quintile | 1.27 (1.14–1.42) |
| Next to highest quintile | 1.22 (1.11–1.33) |
| Middle quintile | 1.23 (1.13–1.34) |
| Next to lowest quintile | 1.25 (1.17–1.34) |
| Lowest quintile | 1.28 (1.19–1.37) |
| Charlson Comorbidity Index | 1.20 (1.10 1.07) |
| ≤1 | 1.42 (1.24–1.63) |
| ≥2 | 1.17 (1.10–1.24) |
| | 1.17 (1.10–1.24) |
| Admissions to hospital in previous year | 1 40 (4 05 4 50) |
| · | 1.46 (1.35–1.58) |
| ≥ 2 | 1.09 (1.01–1.18) |
| Physician clinic visits in previous year | |
| ≤ 6 | 1.32 (1.21–1.43) |
| ≥7 | 1.23 (1.15–1.32) |
| Most responsible diagnosis for readmission to | hospital† |
| Injury | 2.27 (1.72–2.99) |
| Diverticular disease | 2.25 (1.43–3.55) |
| Intestinal obstruction | 1.80 (1.32–2.45) |
| Abdominal pain | 1.61 (0.61–4.23) |
| Skin infection | 1.53 (1.00–2.35) |
| Acute coronary syndrome | 0.93 (0.80–1.07) |
| Venous thromboembolism | 0.89 (0.67–1.19) |
| Chest pain not otherwise specified | 0.75 (0.29–1.90) |
| Syncope and collapse | 0.64 (0.25–1.63) |
| Chronic ischemic heart disease | 0.52 (0.35–0.76) |
| Primary length of stay longer than cohort medi | ian |
| No | 1.26 (1.16–1.36) |
| Yes | 1.25 (1.15–1.35) |
| Hospital-free interval longer than cohort media | an |
| | 1.30 (1.19–1.42) |
| No | () |
| Yes | 1.20 (1.13–1.27) |
| | 1.20 (1.13–1.27) |
| Yes | 1.20 (1.13–1.27) |
| Yes Residence in chronic care facility at readmission | 1.20 (1.13–1.27) on |
| Yes Residence in chronic care facility at readmission No Yes | 1.20 (1.13–1.27) on 1.25 (1.17–1.33) |
| Yes Residence in chronic care facility at readmission No | 1.20 (1.13–1.27) on 1.25 (1.17–1.33) 0.99 (0.89–1.11) |
| Yes Residence in chronic care facility at readmission No Yes Arrival by ambulance at readmission | 1.20 (1.13–1.27) on 1.25 (1.17–1.33) |

Note: CI = confidence interval, OR = odds ratio.

*Calculated for each stratum using unadjusted univariable general estimating equations models.

†The 5 diagnostic categories with the greatest effect on the unadjusted risk of death within 30 days after readmission to an alternative hospital and the 5 with the least effect are presented here. For the complete list of unadjusted results stratified by diagnosis at secondary admission, see eTable 10 in Appendix 1.



Limitations

Our study has several limitations inherent in an observational design. Detailed information, such as clinical history, findings on physical examination, laboratory data, patient preferences, cause of death and indicators of hospital performance, were not available. We had no information on the factors motivating alternative-hospital readmission (e.g., dissatisfaction with prior care, geographic proximity when illness recurred, specialist availability) and could not assess the modifiability of readmission patterns. Information about features that might independently influence a patient's emergency department selection and their health outcomes (e.g., homelessness, frequency of motor vehicle travel and awareness of institutional expertise) was not available. Our focus on a single region also limits the generalizability of the results. The mechanism by which alternative-hospital readmission might increase the risk of death remains a matter of conjecture.

Conclusion

Among patients readmitted within 30 days after discharge, readmission to an alternative hospital was associated with a higher risk of death than readmission to the original hospital. Our study has several implications for patients, clinicians and policy-makers. Before discharge, patients might be informed of the potential advantages of returning to their original hospital should they relapse and require emergency care. Primary care clinicians and prehospital medical service providers might wish to inquire about recent admissions and direct patients back to their original hospital where feasible. Health care financing policies that pay hospitals partly based on observed readmission rates should consider the large numbers of patients readmitted to alternative hospitals and the distinctly high burden of illness associated with such readmissions. In the interim, the data suggest that readmissions to alternative hospitals are worrisome. Whether this adverse prognosis reflects a true causal relation or residual confounding requires further investigation.

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